

SCIENTIFIC REPORTS



OPEN

Emergence of event cascades in inhomogeneous networks

Tomokatsu Onaga & Shigeru Shinomoto

Received: 05 April 2016
 Accepted: 24 August 2016
 Published: 14 September 2016

There is a commonality among contagious diseases, tweets, and neuronal firings that past events facilitate the future occurrence of events. The spread of events has been extensively studied such that the systems exhibit catastrophic chain reactions if the interaction represented by the ratio of reproduction exceeds unity; however, their subthreshold states are not fully understood. Here, we report that these systems are possessed by nonstationary cascades of event-occurrences already in the subthreshold regime. Event cascades can be harmful in some contexts, when the peak-demand causes vaccine shortages, heavy traffic on communication lines, but may be beneficial in other contexts, such that spontaneous activity in neural networks may be used to generate motion or store memory. Thus it is important to comprehend the mechanism by which such cascades appear, and consider controlling a system to tame or facilitate fluctuations in the event-occurrences. The critical interaction for the emergence of cascades depends greatly on the network structure in which individuals are connected. We demonstrate that we can predict whether cascades may emerge, given information about the interactions between individuals. Furthermore, we develop a method of reallocating connections among individuals so that event cascades may be either impeded or impelled in a network.

Our life is full of cause-and-effect relationships, such that past events influence the future occurrence of events. The proliferation process has been studied using both macroscopic models, such as the epidemic model¹, and microscopic models, such as the self-exciting point process proposed by Hawkes^{2,3}. These models have been applied to analyze not only the communication of diseases^{4–8} but also urban crime⁹, human activity^{10–14}, economics¹⁵, genome sequences¹⁶, and neuronal firing^{17,18}. A key quantity representing the interaction in these various phenomena is the basic reproduction ratio, which is defined as the average number of additional events induced by a single event¹⁹. In epidemics, a disease becomes a pandemic in a homogeneous network if the reproduction ratio is greater than unity, as in a nuclear chain reaction^{20–22}, and vanishes otherwise.

Nevertheless, the event-occurrence does not cease if individuals are stimulated in external communities or exhibit spontaneous activity. In such situations, the system may still exhibit cascades of event-occurrences intermittently, even if the reproduction ratio is smaller than the epidemic threshold, as in tweets^{11,23,24} and neuronal firings *in vivo*²⁵. The nonstationary fluctuations may be terminated by reducing the reproduction ratio further²⁶. Event cascades can be a nuisance in some contexts, such as when the peak-demand causes vaccine shortages^{27,28} or heavy traffic on communication lines²⁹, but may be beneficial in other contexts; for example, spontaneous activity in neural networks may be used to generate motion or store memory^{30–32}. Thus, it is important to comprehend the mechanism by which such cascades appear. We show that such a transition between stationary and nonstationary states generally occurs in every proliferation system, obtain the condition on which cascades may emerge in a given network, and suggest a systematic method for controlling systems to oppress or promote the event-occurrence bursts.

Results

Mean rate of event-occurrence. Although the epidemic model and the Hawkes process appear to differ from one another, they have something in common because both were constructed to describe the proliferation processes. To identify their common features, we first revise them by considering realistic constraints.

For an epidemic model, we consider the susceptible-infected-susceptible (SIS) model describing the situation in which infected individuals may recover without immunity:

$$di/dt = \beta si - \gamma i, \quad (1)$$

Department of Physics, Kyoto University, Kyoto 606-8502, Japan. Correspondence and requests for materials should be addressed to S.S. (email: shinomoto@scphys.kyoto-u.ac.jp)

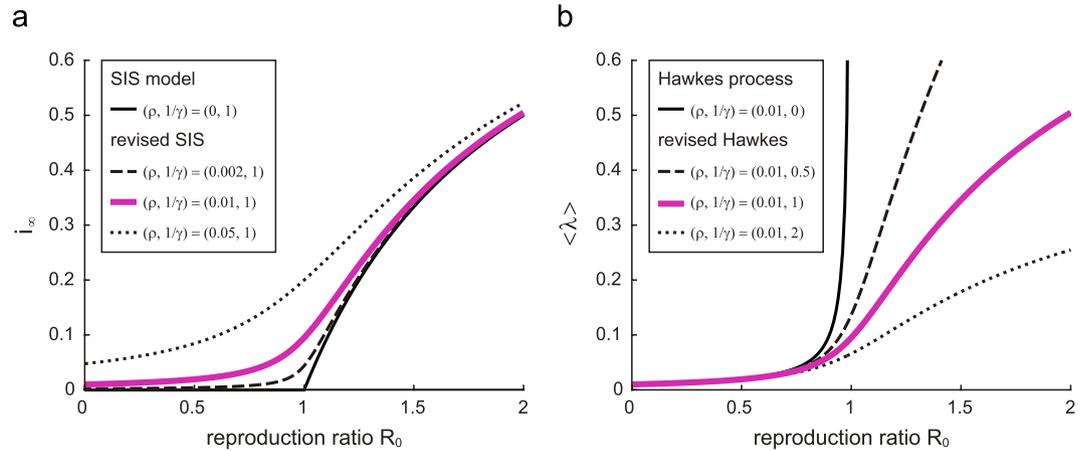


Figure 1. Mean occurrence rate obtained by the revised epidemic model and the revised Hawkes process. (a) The equilibrium fraction of infected individuals i_∞ obtained by the susceptible-infected-susceptible (SIS) model revised by considering spontaneous activity $\rho > 0$. (b) The mean occurrence rate $\langle \lambda \rangle$ of the Hawkes process revised by introducing the refractory period $1/\gamma > 0$. These models give identical equilibria, $\langle \lambda \rangle = \gamma i_\infty$, for the same spontaneous activation ρ and the refractory period $1/\gamma$ (magenta lines in (a,b)).

where i and s are the fractions of infected and susceptible individuals, respectively ($i + s = 1$); β is the rate at which susceptible individuals are infected by contacting infected individuals; and γ is the rate at which infected individuals recover and regain susceptibility. The infected individuals asymptotically vanish if the reproduction ratio $R_0 = \beta/\gamma$ is smaller than or equal to unity; otherwise, the fraction is finite: $i_\infty \equiv \lim_{t \rightarrow \infty} i(t) = 1 - 1/R_0$. To consider extrinsic or spontaneous activation, we suggest revising the SIS model by adding an inflow to the infected population from the susceptible one:

$$di/dt = \beta si - \gamma i + \rho s. \quad (2)$$

In the presence of spontaneous activity $\rho (> 0)$, the asymptotic fraction of infected individuals remains positive and smoothly increases with R_0 , and, accordingly, the epidemic transition at $R_0 = 1$ is softened (Fig. 1a).

The Hawkes process considers spontaneous occurrences in terms of the positive base rate ρ and describes the manner in which the event-occurrence rate $\lambda(t)$ is modulated by past events:

$$\lambda(t) = \rho + R_0 \sum_k h(t - t_k), \quad (3)$$

where t_k is the occurrence time of the k th event. The history kernel $h(t)$ satisfies the causality, $h(t) = 0$ for $t < 0$, and the normalization, $\int_0^\infty h(t) dt = 1$. By taking the ensemble average, the average rate of event-occurrence is obtained as $\langle \lambda(t) \rangle = \rho/(1 - R_0)$. The divergence at $R_0 = 1$ arises from instantaneous reactivation, which is an artifact caused by the simplicity of the linear model, which ignores the refractory period during which each individual does not recover susceptibility. Here, we suggest revising the model by introducing the effect of a refractory period $1/\gamma$:

$$\lambda(t) = \left(1 - \frac{\lambda(t)}{\gamma} \right) \left(\rho + R_0 \sum_k h(t - t_k) \right). \quad (4)$$

By taking an ensemble average and approximating $\langle \lambda^2 \rangle$ with $\langle \lambda \rangle^2$, we obtain the average rate of event-occurrences as $\langle \lambda \rangle = \gamma i_\infty$, where i_∞ is the asymptotic fraction obtained for the revised SIS model (Fig. 1b). Thus, the epidemic model and the Hawkes process may represent the identical mean occurrence rate by taking the spontaneous activation and the refractory periods into account.

Fluctuations in the event-occurrence. Although the systems no longer exhibit a clear transition at $R_0 = 1$ causing the catastrophic chain reaction, they may still show nonstationary fluctuations with intermittent cascades of event-occurrences at lower reproduction ratios; when the reproduction ratio is even smaller, they may remain stationary, producing apparently random events over time.

The SIS model that addresses the mean population dynamics cannot represent fluctuations in the event-occurrences. Here, we construct and simulate a Markov process of microscopic dynamics in which individuals become infected and recover to be susceptible (Fig. 2a). In every interval of a small time-step δt , each susceptible individual in total population of size N may become infected with a probability $(\beta i + \rho)\delta t$, which

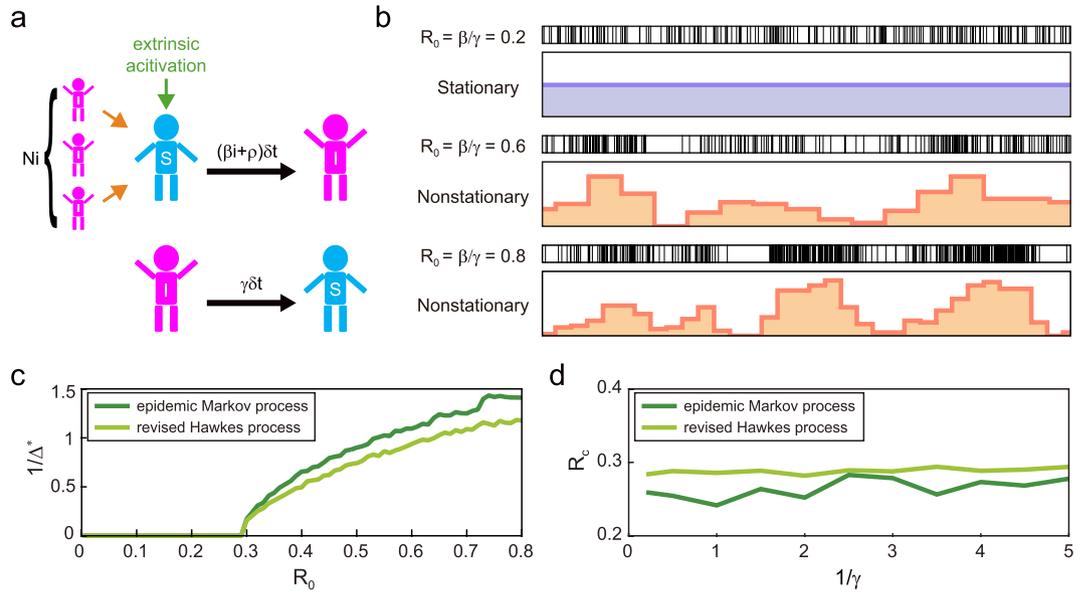


Figure 2. Stationary-nonstationary (SN) transition in the epidemic Markov process and the revised Hawkes process. (a) Microscopic epidemic Markov process. (b) Sample sequences of infected times obtained by the epidemic Markov processes (model parameters: $\beta = 0.06, 0.18,$ and $0.24,$ with $\gamma = 0.3, \rho = 0.5,$ and $N = 1000$). Below the raster diagrams are the fitted optimal histograms. (c) Inverse optimal binsize $1/\Delta^*$ plotted against the reproduction ratio R_0 ($\gamma = 0.3, \rho = 0.5,$ and $N = 1000$). (d) Critical points of the SN transition R_c for the epidemic Markov and revised Hawkes processes, plotted against the refractory period $1/\gamma$.

shifts as $i \rightarrow i + 1/N$. Each infected individual may regain their susceptibility with the probability $\gamma \delta t$, which shifts as $i \rightarrow i - 1/N$. Figure 2b depicts sample sequences of infected times: At a reproduction ratio of $R_0 = \beta/\gamma = 0.2$, the occurrence of infection appears random across individuals but stationary in the whole system. By contrast, at $R_0 = 0.6$ or 0.8 , the event sequence appears nonstationary, exhibiting spontaneous cascades of occurrences. Indeed, microscopic fluctuations are amplified to be macroscopically visible, whereas the average rate is restrained to be finite.

We suggest deciding whether a given series of events is stationary or nonstationary based on whether proper rate estimators conclude a constant rate or a fluctuating rate, respectively. Here, we adopt a method of selecting the histogram bin size to minimize the expected mean square error between the histogram and the unknown underlying rate³³. Note that the decision regarding stationary vs nonstationary state is common across proper rate estimators, such as the Empirical Bayes and variational Bayes Hidden Markov estimators³⁴.

Histograms were fitted to the data: The optimal bin size Δ^* was diverging (becoming as large as the entire observation period) when $R_0 = 0.2$, whereas it was finite (significantly smaller than the entire period) when $R_0 = 0.6$ and 0.8 . Figure 2c depicts the manner in which the inverse bin size $1/\Delta^*$ varies with the reproduction ratio R_0 . $1/\Delta^*$ remains near zero if R_0 is smaller than approximately 0.3 , and it departs from zero otherwise, thus exhibiting the stationary-nonstationary (SN) transition.

We also simulated the revised Hawkes process (4) with parameters identical to those used for the epidemic Markov process. Here, events are indicated by repeating the Bernoulli trials with a probability of $\lambda(t)\delta t$ in every small interval of δt . By plotting $1/\Delta^*$ versus R_0 , we can also observe the transition (Fig. 2c). Figure 2d depicts the critical reproduction ratios obtained for the epidemic Markov process and the revised Hawkes process plotted against the refractory period $1/\gamma$, indicating that the SN critical points for both models are robustly close to $1 - 1/\sqrt{2} \approx 0.3$, which was obtained for the original (linear) Hawkes process ($1/\gamma = 0$) in our previous study²⁶. Note that the critical point is independent of the shape of the kernel, as well as the base rate.

Event-occurrences in inhomogeneous networks. Finally, we consider the emergence of cascades in a population of individuals interacting through inhomogeneous connections.

To obtain the condition for the SN transition analytically, we analyze the linear multivariate Hawkes processes (Fig. 3a). The rate of event-occurrences in the i th individual or node ($i = 1, 2, \dots, N$) is given as

$$\lambda_i(t) = \rho_i + \sum_{j=1}^N \alpha_{ij} \sum_k h(t - t_j^k), \tag{5}$$

where ρ_i represents the base rate, t_j^k is the occurrence time of the k th event in the j th node, and α_{ij} represents the interaction from the j th node to the i th node. Because of the interactions between individuals, $\mathbf{A} \equiv \{\alpha_{ij}\}$, the average rates $\langle \lambda \rangle = \{\langle \lambda_i \rangle\}$ are shifted from the base rates $\rho = \{\rho_i\}$ as $\langle \lambda \rangle = \mathbf{L}\rho$, where \mathbf{L} is the Leontief inverse³⁵: $\mathbf{L} \equiv \sum_{n=0}^{\infty} \mathbf{A}^n = \mathbf{I}/(\mathbf{I} - \mathbf{A})$.

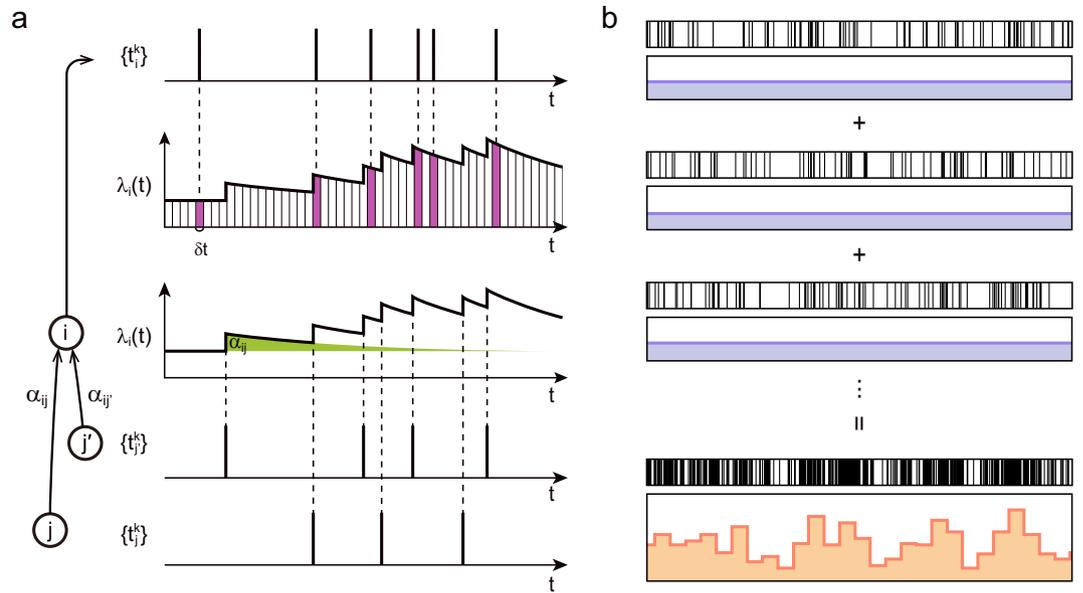


Figure 3. Multivariate Hawkes process. (a) The rate of event-occurrences in each node is modulated by the influence of events generated at other nodes, and events are derived from the underlying rate $\lambda_i(t)$. (b) The manner in which the nonstationary fluctuations become visible by superposing event series in individual nodes.

Even when the fluctuations are not detectable at any single node, all events occurring in the entire set of nodes may exhibit visible fluctuations because the signal-to-noise ratio may increase when multiple series of events are superposed (Fig. 3b). By analyzing equation (5), we obtain the condition for a series of events occurring in an entire population to be nonstationary, or cascading (see Methods):

$$C \equiv \frac{\sum_{i,j} (\mathbf{L}\mathbf{L}^T)_{ij}}{\sum_i \langle \lambda_i \rangle} > 2, \tag{6}$$

where $\mathbf{L} \equiv \text{diag}(\langle \lambda \rangle)$.

Here, we consider 0–1 connectivity with strength $\alpha_{ij} = 0$ or R_0/Nc , where c is the fraction of connections. For a fully connected uniform network ($c = 1$), the summed rate $\lambda(t) = \sum_{i=1}^N \lambda_i(t)$ obeys the original Hawkes process equation (3) with $\rho = \sum_{i=1}^N \rho_i$, thus the entire system may exhibit the SN transition at $R_c = 1 - 1/\sqrt{2} \approx 0.3$. For sparsely connected networks in which every pair of nodes is randomly linked at a fraction of c , the SN critical point R_c remains near 0.3.

Nevertheless, it is possible to shift the cascading condition C by reallocating the connections between individuals by conserving the fraction of connections c ; exchanging connections α_{ij} and $\alpha_{i'j'}$ may alter the cascading condition C by

$$\Delta C = - \frac{\alpha_{ij} H_{ij} - \alpha_{i'j'} H_{i'j'}}{\sum_k \langle \lambda_k \rangle} + \mathcal{O}(R_0/N)^2, \tag{7}$$

where $H_{ij} \equiv ((\sum_k L_{ki})^2 - C)\rho_j + 2(\sum_k L_{kj})\langle \lambda_j \rangle$. To raise or lower C , we repeat exchanging a pair of connections that maximizes or minimizes ΔC (Fig. 4a).

A network may change the state from stationary to nonstationary if the cascading condition C steps across the critical value of 2 from below or vice versa by reconnecting individuals. Figure 4b demonstrates the manner in which C is altered by the steepest ascent or descent based on equation (7). When $(R_0, c) = (0.1, 0.1)$, C remains below 2, even when all connections are reallocated; however, when $(R_0, c) = (0.35, 0.1)$, C exceeds 2, indicating that the system may change between nonstationary ($C > 2$) and stationary ($C \leq 2$).

In the study of epidemics, whether the epidemic threshold is higher in the clustered networks^{8,36,37} or not³⁸ has been controversial. Here, we are not addressing the epidemic transition, but we are interested in how the clustering of individuals influences the SN transition. Figure 4c depicts the manner in which the average clustering coefficient changes with our gradient ascent or descent of C , indicating that clustering tends to facilitate the event-occurrence cascades. A similar tendency was reported in neural network simulations³². An advantage of our method is that we can control the cascade bursting activity by systematically rearranging connections based on a single measure: C .

Figure 4d depicts the range of (R_0, c) in which a network of a large size N may exhibit stationary and nonstationary states. In the “Stationary” regime, the systems never generate visible cascades, even when connections are reallocated, whereas in the “Nonstationary” regime, the systems always exhibit cascades. In the “S-N 0” regime, networks may be either stationary or nonstationary, depending on the manner in which individuals are connected. In the “S-N 1” regime, networks may be either stationary or nonstationary if the reciprocal connections

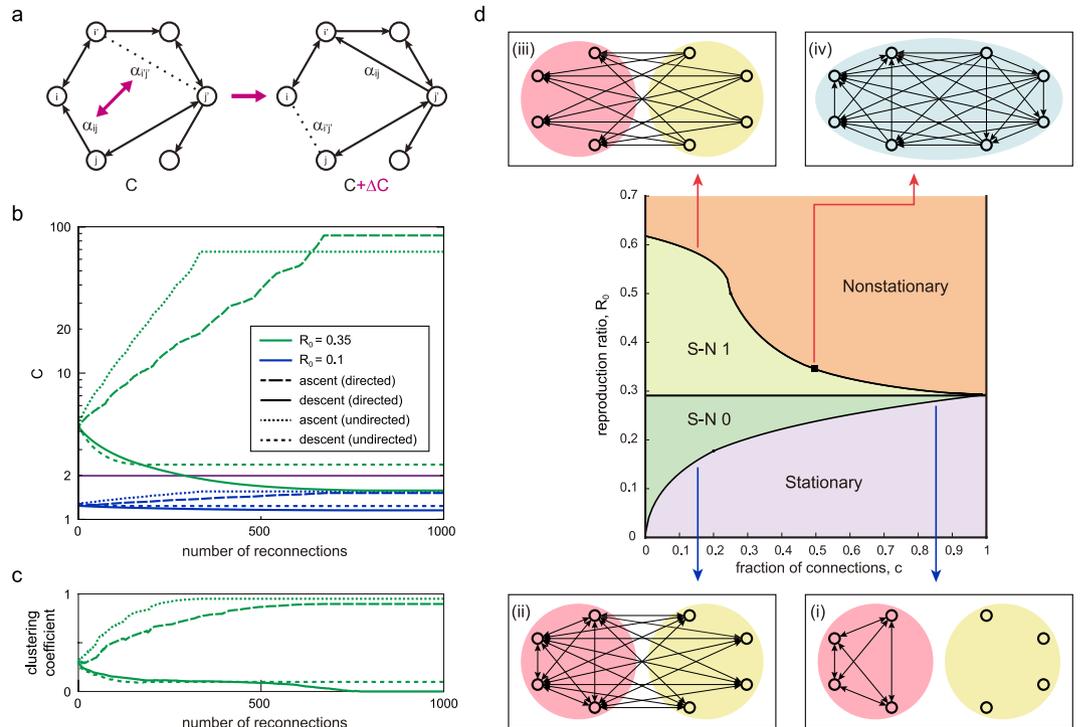


Figure 4. Controlling the emergence of event-occurrence cascades. (a) Elementary process of exchanging connections α_{ij} and α_{ji} . (b) The manner in which the potential for the cascades C is altered by the steepest ascent or descent based on equation (7), starting from the Erdős-Rényi models. The blue and green lines represent the cases of $(R_0, c) = (0.1, 0.1)$, and $(0.35, 0.1)$, respectively ($N = 100$). (c) Changes in the average clustering coefficient according to the reallocation of connections when $(R_0, c) = (0.35, 0.1)$. (d) Parameter ranges of (R_0, c) in which the networks may be either stationary or nonstationary. Some solvable extreme configurations that give low and high critical points in R_0 are depicted.

can be controlled independently, whereas undirected networks, whose connections are reciprocally symmetric $\{\alpha_{ij} = \alpha_{ji}\}$, remain nonstationary, always generating cascades of events. For undirected networks, the critical point R_c appears to be bounded at approximately 0.3.

As extreme configurations with low critical points, we considered networks of (i): the nodes of one group are fully connected within the group, and the others are isolated; and (ii): the nodes of one group are fully connected within the group and also receive undirected (reciprocal) connections from all nodes of another group. The critical points R_c exhibit crossover at $c = 0.20$, and above and below this point, those of (i) and (ii) are lower. The criticality in low R_0 implies that these configurations tend to facilitate cascade bursting. As an opposite extreme, we considered another configuration (iii): each node of one group receives only directed connections from another group. We also considered a specific hierarchical networks of (iv): every node exerts the influence over the lower hierarchy nodes in a manner that connections form a triangular matrix. All critical points for these specific cases are obtained analytically (see Methods) and plotted in Fig. 4d: The configurations (i) and (ii) tend to incite bursting, whereas the configurations (iii) and (iv) impede bursting. These observations imply that networks in which small number of individuals occupy reciprocal connections favor event cascade bursting, whereas directed unilateral connections tend to impede cascades.

Any network exhibits the SN transition within the range of S-N 0 or S-N 1 on a plane of (R_0, c) , independently of the system size N (Fig. 4d). The precise transition point within the transition range is determined by the network configuration $A \equiv \{\alpha_{ij}\}$. Here we examine the SN transition in representative network models, such as the scale-free network (Barabási-Albert model)³⁹, the random network (Erdős-Rényi model)⁴⁰, and the small-world network (Watts-Strogatz model)⁴¹ assuming reciprocal connections. In particular, we compare the SN critical points R_c of those networks of identical number of nodes N and the fraction of connections inversely proportional to the number of nodes $c = 6/N$ (Fig. 5). The scale-free network is assembled with the Barabási-Albert model. The small-world network is made by rewiring a regular ring lattice with the rewiring probability $p = 0.1$. The critical point R_c in the scale-free network is the lowest among those networks, implying that the wide disparity in degrees (numbers of connections) of individual nodes may facilitate the fluctuations in event occurrences. By obtaining the average critical point analytically, we may confirm that $R_c \rightarrow 0$ as $N \rightarrow \infty$ (see Methods). It is interesting to see that R_c is higher in the small-world network than in the Erdős-Rényi model. This may also be closely related to the difference in degrees; the small-world network of the small rewiring probability exhibits the smaller degree dispersion than the Erdős-Rényi model. Indeed, a regular ring lattice in which individual degrees are identical gives the highest critical point $R_c = 0.2929$, which is equal to the critical point of a fully connected uniform network,

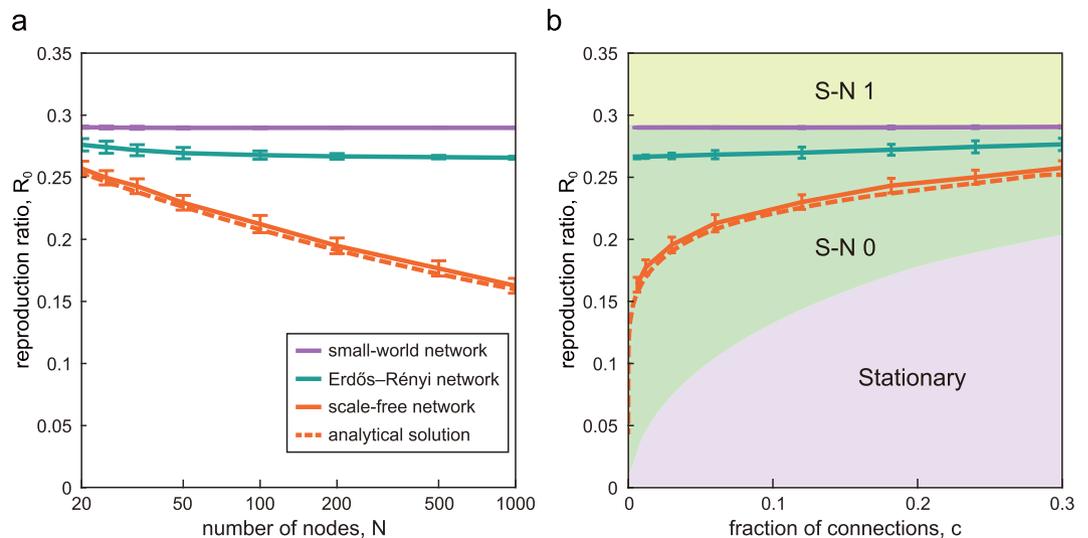


Figure 5. Stationary-nonstationary (SN) transitions in representative complex networks. (a) The manner in which the critical reproduction ratio R_c varies with the number of nodes N in the scale-free network (Barabási-Albert model), the random network (Erdős-Rényi model), and the small-world network (Watts-Strogatz model), possessing reciprocal connections. In all networks, the fraction of connections varied inversely proportional to the system size as $c = 6/N$. (b) Dependence of R_c on c plotted on a plane in Fig. 4d. The error bars represent standard deviation for 100 sample networks. The dashed line represents the average critical point analytically obtained for the scale-free network.

$1 - 1/\sqrt{2}$. In this way, the SN transition points R_c can be different, crucially depending on the network structure, in particular on the degree dispersion.

Finally we demonstrate the manner in which the network is reconnected by our method for inciting or impeding cascades (Fig. 6); we take up the network of “Zachary’s Karate Club”⁴² and reallocate the (undirected) friendships between 34 people. By assuming that the original network is at the criticality $C = 2$, we reorganized the network so that their entire communication becomes stationary or nonstationary. The operation to impede the communication cascades ($\Delta C < 0$) tended to make all people be connected harmoniously. Contrariwise, the operation to incite the cascade ($\Delta C > 0$) made the society polarized into two kinds of people; a few people have exclusive rich connections, while others are connected only with those people.

Discussion

Here, we showed that the proliferation process may exhibit the SN transition at which nonstationary cascades of event-occurrences emerge from the stationary process. The critical reproduction ratio for the SN transition R_c is much smaller than the threshold for the conventional epidemic transition exhibiting chain reactions. The SN transition may depend greatly on the manner in which individuals are connected. We developed a theory for predicting the occurrence of cascades in a network. We also suggested a method of reconnecting individuals for impeding or inciting cascade bursting in a network. It was revealed that networks of reciprocal connections exhibiting wide disparity in degrees of individual nodes are apt to trigger the event cascades.

Note that the term “cascade” has also been used differently in other contexts, as expressing transient phenomena of event spreads, as in failures in a power grid, chain bankruptcies, and information cascades in social systems^{43–45}. Though their basic process is similar to that of our SN transition, the situations are different; we considered here the fluctuations in the process of event occurrences that involves constant spontaneous activity or external activation such that individuals are continually generating events at a finite probability.

In this study, we revealed that nonstationary fluctuations may occur autonomously in a static network without environmental fluctuations. In real-world circumstances, however, both connections and environmental inputs may change in time, and the event cascades may have been induced by such extrinsic cause. Thus it may be an important direction in future to consider discerning the factor that may have contributed to the observed fluctuations, namely, whether the fluctuations were originated intrinsically or extrinsically.

To set about applying our theory to real-world problems, we should also consider adapting models to individual situations, by collecting information regarding interactions among individuals. Interactions in epidemic networks may be inferred by getting hold of information of contacts between individuals and estimating the infectivity of the disease. Interactions in neural networks may be estimated by detecting the synaptic connections and measuring the strength of connections such as excitatory or inhibitory postsynaptic potentials. If the information of interactions is given, we can predict whether the network may exhibit nonstationary fluctuations, using the theory developed here. However, it is practically difficult to obtain the entire information of the interactions among all pairs of existing nodes. Thus it is useful to devise a method of complementing information of the entire network from partial observation, and furthermore, to suggest controlling systems by manipulating partially available links.

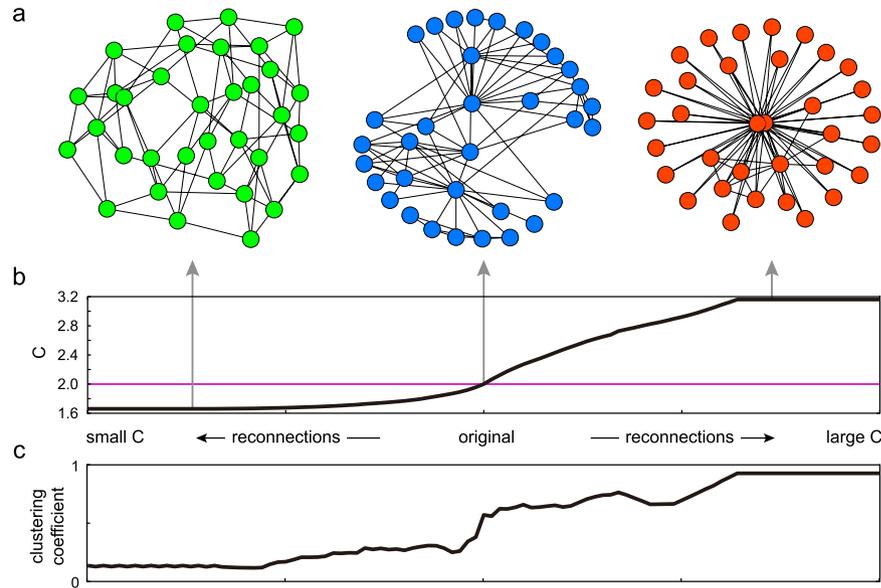


Figure 6. Reallocating connections between individuals. (a) The manner in which the (undirected) friendships between 34 people in “Zachary’s Karate Club” are reconnected by increasing or decreasing C , respectively for inciting or impeding cascades of communications. (b) The change in the potential for the cascades C with the steepest ascent (rightward) or descent (leftward) based on equation (7), (c) Average clustering coefficient.

Methods

Derivation of the condition for the SN transition. Here, we derive the condition for the SN transition for a given series of events. It has been proven that the optimal bin size may be finite if fluctuation in the underlying rate $\delta\lambda(t) \equiv \lambda(t) - \langle \lambda \rangle$ satisfies the condition^{34,46}

$$\frac{1}{\langle \lambda \rangle} \int_{-\infty}^{\infty} \langle \delta\lambda(t+s)\delta\lambda(t) \rangle ds > 1, \tag{8}$$

and diverges otherwise. This condition (8) is derived as follows. The mean square error between the underlying rate $\lambda(t)$ and the histogram $\hat{\lambda}(t)$ is given as

$$S = \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T \langle (\lambda(t) - \hat{\lambda}(t))^2 \rangle dt, \tag{9}$$

where T is the entire observation interval, and the bracket represents the ensemble average over the possible realization of the stochastic process. In each bin of size Δ , the histogram $\hat{\lambda}(t)$ is a constant whose height is the number of events K divided by the bin size Δ . Thus, the mean square error is transformed as

$$S = \left\langle \frac{1}{\Delta} \int_0^\Delta \left(\lambda^2(t) - \frac{2K}{\Delta} \lambda(t) + \frac{K^2}{\Delta^2} \right) dt \right\rangle. \tag{10}$$

The expected number of events in each interval is given by integrating the underlying rate: $\langle K \rangle = \int_0^\Delta \lambda(t) dt$. Because events are independently drawn, the Poisson relation holds: $\langle K^2 \rangle = \langle K \rangle^2 + \langle K \rangle$. Inserting these relations into equation (10), we have

$$S = \phi(0) + \frac{\langle \lambda \rangle}{\Delta} - \frac{1}{\Delta^2} \int_0^\Delta dt \int_{-t}^t \phi(s) ds, \tag{11}$$

where $\phi(s) \equiv \langle \lambda(t+s)\lambda(t) \rangle - \langle \lambda \rangle^2$ is the correlation of the rate fluctuation or $\phi(s) = \langle \delta\lambda(t+s)\delta\lambda(t) \rangle$, where $\delta\lambda(t) \equiv \lambda(t) - \langle \lambda \rangle$ is the temporal fluctuation of the rate. The mean square error may have a minimum at some finite Δ . Based on the second-order transition in which the minimum position Δ^* goes to infinity or $1/\Delta^*$ goes to zero continuously, the condition for the transition is given as

$$\left. \frac{dS}{d(1/\Delta)} \right|_{\Delta=\infty} < 0. \tag{12}$$

This can be summed up as a condition of the rate fluctuation given in inequality (8) if $\int_0^\infty s\phi(s) ds$ is finite. This condition was derived from the optimization of a histogram and was found to be identical to that derived from the marginal likelihood maximization of the Bayesian rate estimator, implying that this condition may be a universal bound for detecting rate fluctuation⁴⁷.

For the linear self-exciting point process, the power spectrum of the rate fluctuation or the Fourier transformation of the autocorrelation $\phi(s) \equiv \langle \delta\lambda(t+s)\delta\lambda(t) \rangle$ was obtained by Hawkes². The result can be summarized as²⁶

$$\tilde{\phi}_\omega = \left(\frac{1}{(1 - R_0\tilde{h}_\omega)(1 - R_0\tilde{h}_{-\omega})} - 1 \right) \langle \lambda \rangle, \tag{13}$$

where \tilde{h}_ω is the Fourier transform of the kernel function $h(t)$. Because $\tilde{\phi}_0 = \int_{-\infty}^{\infty} \phi(s) ds = \int_{-\infty}^{\infty} \langle \delta\lambda(t+s)\delta\lambda(t) \rangle ds$, the condition for the linear self-exciting process to be nonstationary is obtained as $1/(1 - R_0)^2 > 2$. Thus, the SN transition occurs at $R_0 = 1 - 1/\sqrt{2}$ independent of the time course of supplementary probability $h(t)$ and the base rate ρ .

The multivariate Hawkes process (5) is also analytically tractable; in particular, the Fourier zero-mode of the correlations $\phi(s) \equiv \{\phi_{ij}(s)\} \equiv \{\langle \delta\lambda_i(t+s)\delta\lambda_j(t) \rangle\}$ is obtained as^{3,26},

$$\tilde{\phi}_0 = \mathbf{L}\mathbf{A}\mathbf{L}^T - \mathbf{A}. \tag{14}$$

Each node may exhibit the SN criticality if the correlation of each individual, $\phi_{ii}(s)$, satisfies the SN condition. Even when rate fluctuations are not detectable at any single node, the summed activity of multiple nodes may exhibit fluctuations. The condition for the superposed series to exhibit the SN transition is obtained by applying the nonstationary condition (8) to the summed rate $\lambda(t) = \sum_{i=1}^N \lambda_i(t)$, thus leading to the cascading condition, equation (6).

Criticality conditions for extreme configurations. In the following, we give the criticality conditions $C=2$ in equation (6) for the proposed extreme configurations (i), (ii), (iii), and (iv) (Fig. 4d). The connectivity of configuration (i) is given as

$$A = \frac{R_0}{Nc} \begin{pmatrix} \mathbf{1}_{M,M} & \mathbf{0}_{M,N-M} \\ \mathbf{0}_{N-M,M} & \mathbf{0}_{N-M,N-M} \end{pmatrix}, \tag{15}$$

where $\mathbf{1}_{n,m}$ and $\mathbf{0}_{n,m}$ are $n \times m$ matrices consisting of all elements of 1 and 0, respectively. In this case, the fraction of connections is related to M as $c = M^2/N^2$, and the critical point is obtained by solving a cubic equation:

$$(x - c^{1/2})^3 = c^{1/2}x^3 - cx^2 - c^{3/2}x. \tag{16}$$

The connectivity of the configuration (ii) is given as

$$A = \frac{R_0}{Nc} \begin{pmatrix} \mathbf{1}_{M,M} & \mathbf{1}_{M,N-M} \\ \mathbf{1}_{N-M,M} & \mathbf{0}_{N-M,N-M} \end{pmatrix}. \tag{17}$$

The critical point is obtained by solving a fifth-degree equation:

$$\begin{aligned} & 2b^3x^5 + 2b^2(3 - 2b)(1 + b)x^4 + b(6 - 12b + 5b^2)(1 + b)^2x^3 \\ & + (1 - b)(2 - 10b + b^2)(1 + b)^3x^2 - (1 - b)^2(4 + b)(1 + b)^4x \\ & + (1 - b)^2(1 + b)^5 = 0, \end{aligned} \tag{18}$$

where $b = \sqrt{1 - c} = (N - M)/N$.

The connectivity of the configuration (iii) is given as

$$A = \frac{R_0}{Nc} \begin{pmatrix} \mathbf{0}_{N-M,N-M} & \mathbf{0}_{N-M,M} \\ \mathbf{1}_{M,N-M} & \mathbf{0}_{M,M} \end{pmatrix}. \tag{19}$$

In this case, c is related to M with $c = M(N - M)/N^2$, and the critical point is obtained as

$$R_0 = \frac{1 + \sqrt{1 - 4c}}{4} (-1 + \sqrt{1 + 2c^{-1}(1 - \sqrt{1 - 4c})}). \tag{20}$$

The connectivity of the configuration (iv) is given by the triangular matrix,

$$A = \frac{R_0}{Nc} \begin{pmatrix} 0 & 0 & \dots & 0 \\ 1 & \ddots & \ddots & \vdots \\ \vdots & \ddots & \ddots & 0 \\ 1 & \dots & 1 & 0 \end{pmatrix}. \tag{21}$$

In this case, $c = \lim_{N \rightarrow \infty} (N - 1)/(2N) = 1/2$, and the critical point is obtained as

$$R_0 = \lim_{N \rightarrow \infty} \frac{N - 1}{2} (2^{1/(N-1)} - 1) = \frac{\log 2}{2} \approx 0.35 \tag{22}$$

Critical reproduction ratio of the scale-free network. Here, we derive the critical value R_c in the scale free network using degree based mean field approximation⁴⁸. The critical point is obtained by solving a cubic equation of $x=R_c$:

$$\begin{aligned} & (\langle k^3 \rangle \langle k \rangle^3 + \langle k^2 \rangle^3 - 2 \langle k^2 \rangle^2 \langle k \rangle^2) x^3 + (\langle k^2 \rangle \langle k \rangle^4 - 3 \langle k^2 \rangle^2 \langle k \rangle^2) x^2 \\ & + (3 \langle k^2 \rangle \langle k \rangle^4 + \langle k \rangle^6) x - \langle k \rangle^6 = 0, \end{aligned} \quad (23)$$

where k is the degree of each node and $\langle \cdot \rangle$ represents the average over the degree distribution $P(k)$. If $P(k) \sim k^{-\gamma} (\gamma \leq 3)$, we may confirm that $R_c \rightarrow 0$ as $N \rightarrow \infty$.

References

- Kermack, W. O. & McKendrick, A. G. A contribution to the mathematical theory of epidemics. *Proc. Roy. Soc. Lond. A* **115**, 700–721 (1927).
- Hawkes, A. G. Spectra of some self-exciting and mutually exciting point processes. *Biometrika* **58**, 83–90 (1971).
- Hawkes, A. G. Point spectra of some mutually exciting point processes. *J. Roy. Statist. Soc. Ser. B* **33**, 438–443 (1971).
- Hethcote, H. W. The mathematics of infectious diseases. *SIAM Review* **42**, 599–653 (2000).
- Keeling, M. J. & Rohani, P. *Modeling infectious diseases in humans and animals* (Princeton University Press, 2008).
- Vespignani, A. Modelling dynamical processes in complex socio-technical systems. *Nat. Phys.* **8**, 32–39 (2012).
- Brockmann, D. & Helbing, D. The hidden geometry of complex, network-driven contagion phenomena. *Science* **342**, 1337–42 (2013).
- Pastor-Satorras, R., Castellano, C., Van Mieghem, P. & Vespignani, A. Epidemic processes in complex networks. *Rev. Mod. Phys.* **87**, 925–979 (2015).
- Mohler, G. O., Short, M. B., Brantingham, P. J., Schoenberg, F. P. & Tita, G. E. Self-Exciting Point Process Modeling of Crime. *J. Am. Stat. Assoc.* **106**, 100–108 (2011).
- Goffman, W. & Newill, V. Generalization of epidemic theory. *Nature* **204**, 225–228 (1964).
- Crane, R. & Sornette, D. Robust dynamic classes revealed by measuring the response function of a social system. *Proc. Natl. Acad. Sci. USA* **105**, 15649–15653 (2008).
- Kitsak, M. *et al.* Identification of influential spreaders in complex networks. *Nat. Phys.* **6**, 888–893 (2010).
- Lerman, K. & Ghosh, R. Information contagion: An empirical study of the spread of news on digg and twitter social networks. *ICWSM* **10**, 90–97 (2010).
- Romero, D. M., Meeder, B. & Kleinberg, J. Differences in the mechanics of information diffusion across topics: idioms, political hashtags, and complex contagion on twitter. *Proc. 20th Intl. WWW Conf.* 695–704 (ACM, 2011).
- At-Sahalia, Y., Cacho-Diaz, J. & Laeven, R. J. Modeling financial contagion using mutually exciting jump processes. *Tech. Rep.*, National Bureau of Economic Research (2010).
- Reynaud-Bouret, P., Schbath, S. *et al.* Adaptive estimation for hawkes processes; application to genome analysis. *Ann. Stat.* **38**, 2781–2822 (2010).
- Pernice, V., Staude, B., Cardanobile, S. & Rotter, S. How structure determines correlations in neuronal networks. *PLoS Comput. Biol.* **7**, e1002059 (2011).
- Reynaud-Bouret, P., Rivoirard, V., Grammont, F. & Tuleau-Malot, C. Goodness-of-fit tests and nonparametric adaptive estimation for spike train analysis. *J. Math. Neurosci.* **4**, 3 (2014).
- Heffernan, J., Smith, R. & Wahl, L. Perspectives on the basic reproductive ratio. *J. R. Soc. Interface* **2**, 281–293 (2005).
- Anderson, H. L., Fermi, E. & Szilard, L. Neutron production and absorption in uranium. *Phys. Rev.* **56**, 284 (1939).
- Scharff-Goldhaber, G. & Klaiber, G. Spontaneous emission of neutrons from uranium. *Phys. Rev.* **70**, 229 (1946).
- Duderstadt, J. J. & Hamilton, L. J. *Nuclear reactor analysis* (1976).
- Barabási, A.-L. The origin of bursts and heavy tails in human dynamics. *Nature* **435**, 207–211 (2005).
- Kwak, H., Lee, C., Park, H. & Moon, S. What is twitter, a social network or a news media? *Proc. 19th Intl. WWW Conf.* 591–600 (ACM, 2010).
- Sakata, S. & Harris, K. D. Laminar structure of spontaneous and sensory-evoked population activity in auditory cortex. *Neuron* **64**, 404–418 (2009).
- Onaga, T. & Shinomoto, S. Bursting transition in a linear self-exciting point process. *Phys. Rev. E* **89**, 042817 (2014).
- Hinman, A. R., Orenstein, W. A., Santoli, J. M., Rodewald, L. E. & Cochi, S. L. Vaccine shortages: History, impact, and prospects for the future. *Annu. Rev. Public Health* **27**, 235–259 (2006).
- Stöhr, K. Vaccinate before the next pandemic? *Nature* **465**, 161 (2010).
- Huberman, B. A. & Lukose, R. M. Social dilemmas and internet congestion. *Science* **277**, 535–537 (1997).
- Jaeger, H. & Haas, H. Harnessing nonlinearity: predicting chaotic systems and saving energy in wireless communication. *Science* **304**, 78–80 (2004).
- Sussillo, D. & Abbott, L. F. Generating coherent patterns of activity from chaotic neural networks. *Neuron* **63**, 544–557 (2009).
- Laje, R. & Buonomano, D. V. Robust timing and motor patterns by taming chaos in recurrent neural networks. *Nat. Neurosci.* **16**, 925–933 (2013).
- Shimazaki, H. & Shinomoto, S. A method for selecting the bin size of a time histogram. *Neural Comput.* **19**, 1503–1527 (2007).
- Shintani, T. & Shinomoto, S. Detection limit for rate fluctuations in inhomogeneous poisson processes. *Phys. Rev. E* **85**, 041139 (2012).
- Waugh, F. V. Inversion of the leontief matrix by power series. *Econometrica* **18**, 142–154 (1950).
- Keeling, M. J. The effects of local spatial structure on epidemiological invasions. *Proc. R. Soc. Lond. B* **266**, 859–867 (1999).
- Miller, J. C. Percolation and epidemics in random clustered networks. *Phys. Rev. E* **80**, 020901 (2009).
- Newman, M. E. Properties of highly clustered networks. *Phys. Rev. E* **68**, 026121 (2003).
- Barabási, A.-L. & Albert, R. Emergence of scaling in random networks. *Science* **286**, 509–512 (1999).
- Erdős, P. & Rényi, A. On random graphs, i. *Publicationes Mathematicae (Debrecen)* **6**, 290–297 (1959).
- Watts, D. J. & Strogatz, S. H. Collective dynamics of ‘small-world’ networks. *Nature* **393**, 440–442 (1998).
- Zachary, W. W. An information flow model for conflict and fission in small groups. *J. Anthropol. Res.* **33**, 452–473 (1977).
- Watts, D. J. A simple model of global cascades on random networks. *Proc. Natl. Acad. Sci. USA* **99**, 5766–5771 (2002).
- Buldyrev, S. V., Parshani, R., Paul, G., Stanley, H. E. & Havlin, S. Catastrophic cascade of failures in interdependent networks. *Nature* **464**, 1025–1028 (2010).
- Zhao, J., Li, D., Sanhedrai, H., Cohen, R. & Havlin, S. Spatio-temporal propagation of cascading overload failures in spatially embedded networks. *Nat. Commun.* **7** (2016).
- Koyama, S. & Shinomoto, S. Histogram bin width selection for time-dependent poisson processes. *J. Phys. A Math. Gen.* **37**, 7255 (2004).

47. Koyama, S., Shimokawa, T. & Shinomoto, S. Phase transitions in the estimation of event rate: a path integral analysis. *J. Phys. A Math. Theor.* **40**, F383 (2007).
48. Pastor-Satorras, R. & Vespignani, A. Epidemic spreading in scale-free networks. *Phys. Rev. Lett.* **86**, 3200–3203 (2001).

Acknowledgements

We thank Takaaki Aoki, Hiroya Nakao, and Naoki Masuda for their advice and comments. This study was supported in part by a Grant-in-Aid for Scientific Research to S.S. from MEXT Japan (26280007) and by JST and CREST. T.O. is supported by a JSPS Research Fellowship for Young Scientists.

Author Contributions

S.S. designed the study. T.O. conducted the theoretical and numerical analysis. S.S. wrote the manuscript.

Additional Information

Competing financial interests: The authors declare no competing financial interests.

How to cite this article: Onaga, T. and Shinomoto, S. Emergence of event cascades in inhomogeneous networks. *Sci. Rep.* **6**, 33321; doi: 10.1038/srep33321 (2016).



This work is licensed under a Creative Commons Attribution 4.0 International License. The images or other third party material in this article are included in the article's Creative Commons license, unless indicated otherwise in the credit line; if the material is not included under the Creative Commons license, users will need to obtain permission from the license holder to reproduce the material. To view a copy of this license, visit <http://creativecommons.org/licenses/by/4.0/>

© The Author(s) 2016