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RATE $O((n^{-1} \log n)^{\frac{1}{2}})$ OF RISK CONVERGENCE IN THE EMPIRICAL BAYES
BOOTSTRAP ESTIMATION: CASE OF RETRACTED DISTRIBUTIONS

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1. Introduction.

The empirical Bayes (EB) problem involved here is the same one as stated in Nogami (1983). We use the same notational conventions. We often let $P(h)$ or $P(h(w))$ denote $\int h(w) dP(w)$.

For the uniform distribution $U[0, \theta+1]$ where $\theta \in (-\infty, \infty)$ Fox (1970) exhibited a distribution-valued Lévy consistent estimate $\hat{G}_n$ for $G$. In the empirical Bayes problem where the $\theta_i$ are iid with a common unknown prior distribution $G$, Fox (1970) indicated (without rates) a convergence of the expected risks to the Bayes envelope $R(G)(=R)$ for a bootstrap decision rule based on component procedures Bayes vs $\hat{G}_n$. In this paper we consider the square error loss estimation and exhibit two distribution-valued Lévy consistent estimates $\hat{G}_n$ and $\hat{G}_n$ for $G$ involving the family $p(f)$ of retracted distributions into the interval $[0, \theta+1]$. We furthermore obtain two bootstrap estimates $\hat{\phi}_n$ and $\hat{\phi}_n$ (for $\theta$) with a convergence rate $O((n^{-1} \log n)^{\frac{1}{2}})$ for both $R(\hat{\phi}_n, G)-R$ and $R(\hat{\phi}_n, G)-R$.

This paper is essentially dealt with an application of the compound estima in Nogami (1982) for the empirical Bayes problem. Differently from Nogami (1982) we here exhibits two kernel type EB estimates and the proof of Lemma 4.5 is slightly simpler than that in Nogami (1982, Proof of Lemma 2.5).
In Section 2 we obtain an upper bound of $R(\hat{\theta}_\mu, G) - R$ for any bootstrap estimate $\hat{\theta}_\mu$ (for $\theta$). In Section 3 we introduce two EB estimates $\hat{\phi}_n$ and $\hat{\phi}_n$. We shall obtain, in Section 4, a rate $O((n^{-1}\log n)^{1/2})$ for $R(\hat{\phi}_n, G) - R$ and, in Section 5, the same rate for $R(\hat{\phi}_n, G) - R$.

2. An Upper Bound for $R(\hat{\theta}_n, G) - R$.

Let $\Omega = (-\infty, \infty)$. Assume $(0<)m^{-1} \leq 1$ throughout the paper. Let $\hat{\xi}_n$ be any distribution-valued random variable which is an estimate of the unknown prior distribution $G$, obtained from $X_1, \ldots, X_n, X_{n+1}$ where each $X_i - P_{\theta}/P_\theta$ dG($\theta$) for $P_\theta \in P(\xi)$ and $X - P_{\theta}(\theta = \theta_{n+1})$. Define an estimate for $\theta$ by

$$
\hat{\theta}_n(X) = \int_{X_1}^{X} q(\theta) \, d\hat{G}_n(\theta)/ \int_{X_1}^{X} q(\theta) \, d\hat{G}_n(\theta).
$$

When $R(\hat{\theta}_n, G)$ and $R$ are finite, we have

$$
(0<) R(\hat{\theta}_n, G) - R = E(\hat{\theta}_n(X) - \phi_G(X))^2
$$

where $\phi_G(X)$ is the Bayes estimate vs $G$ at $X$ of the form

$$
\phi_G(X) = \int_{X_1}^{X} q(\theta) \, dG(\theta)/ \int_{X_1}^{X} q(\theta) \, dG(\theta)
$$

where the affix $+$ means the right limit and hereafter we abbreviate it.

To get a bound we need to set the following assumptions on $G$:

A i) For some $(0<)\epsilon < 1$ and a positive constant $B$

$$
\sup_{s} \sup_{y \in (s, s+\epsilon)} |G'(y)| \leq B (\epsilon, \infty).
$$

A ii) For a positive constant $C$,

$$
\sup_{(0<)\epsilon < 1/3} \int \{G(y+1-\epsilon) - G(y+\epsilon)\}^{-1} \, dy \leq C (\epsilon, \infty).
$$

Lévy distance for two distribution functions $F$ and $H$ of random variables (cf. Feller (1971), p. 285) is defined by

$$
L(F, H) = \inf_{\epsilon > 0: F(y-\epsilon) - \epsilon \leq H(y) \leq F(y+\epsilon) + \epsilon} \text{ for all } y \in (-\infty, \infty).
$$
Remark that the infimum in the definition attains (see Appendix of Nogami (1975)).

In this section we shall find an upper bound for (1.2). Since by (1.1) and (1.3) \( X < \hat{\ell}_n(X), \phi_n(X) \leq X \), we have \( |\hat{\ell}_n - \phi_n| \leq 1 \). Hence, for any \( 0 < \varepsilon < 1 \),

\[
E(\hat{\ell}_n(X) - \phi_n(X))^2 \leq E[L(G, \hat{G}_n) > \varepsilon] + E[(\hat{\ell}_n(X) - \phi_n(X))^2[L(G, \hat{G}_n) \leq \varepsilon]].
\]

The main development of this section is Lemma 1.6 in which we show that two terms in the rhs(1.6) below are both no more than a constant times \( \varepsilon^2 \) for \( 0 < \varepsilon < 1 \) which in turn the second term of rhs(1.4) is no more than a constant times \( \varepsilon^2 \).

To reach the bound in Lemma 1.6 we shall use Lemma A.2 of Singh (1974), Proposition A of Nogami (1975) and Lemma A.3 of Nogami (1975). Hence, we shall state them here beforehand.

**Lemma 1.1.** (R. S. Singh (1974)) Let \( y, z \) and \( B \) be in \((-\infty, \infty)\) with \( z \not= 0 \) and \( B > 0 \). If \( Y \) and \( Z \) are two real valued random variables, then for every \( \gamma > 0 \)

\[
E(Y/Z) - (y/z) \geq a^+ \gamma + (y/z)^\gamma + 2^{-(\gamma-1)^+} B^\gamma E|y - y|^{\gamma} + (y/z)^\gamma + 2^{-(\gamma-1)^+} B^\gamma E|z - z|^{\gamma}
\]

where \( a^+ = a \not= 0 \).

**Lemma 1.2.** (Proposition A of Nogami (1975)) Let \( I \) be a finite interval and let \( F_I \) be the retraction of a distribution function \( F \) into the closed interval \([F(a^+), F(b^+)]\). Then,

\[
L(F_I, G_I) \leq |F-G)(a^+)| \vee (F-G)(b^+) \vee L(F, G)
\]

where we use \( + \) on the line to denote the right limit.

For Lemma 1.3 we need to introduce the following definition:

**Definition 1.1.** With \( h \), a function defined on a real interval \( I \), the modulus of continuity of \( h \) is the function given by

\[
\alpha(\varepsilon) = \sup\{h(\omega_1) - h(\omega_2): \omega_1, \omega_2 \not= I, |\omega_1 - \omega_2| < \varepsilon\}
\]

for every \( \varepsilon > 0 \).
In Lemma 1.3 below, the natural generalization of the inverse probability integral transformation is used to develop bounds for the same difference of integrals in Lemma 8' of Oaten (1969, Appendix) without partitioning as in Oaten's proof.

**Lemma 1.3.** (Nogami (1975, Lemma A.3)) Let $I$ be a finite interval $[a, b]$ supporting finite measures $\mu$ and $\nu$ and let $h$ be measurable on $I$ into a finite interval $[c, d]$. Let $F$ and $G$ be distribution functions inducing $\mu$ and $\nu$ with $F(a-)\nu G(a-) \leq F(b+)\Lambda G(b+).$ Then $\int h \, d(\mu - \nu)$ has the following family of bounds

$$
\frac{d-c}{2} \left| (F-G)(a-) + (F-G)(b+) \right|
+ \alpha L(F,G)+(F(b)+)G(b+)-F(a-)VG(a-)) + \frac{d+c}{2} |\mu(I) - \nu(I)|.
$$

We first apply Lemma 1.1 of Nogami (1981) for $\hat{\phi}_n(X)$ and $\phi_G(X)$. (This lemma is essentially the Fubini Theorem.) Since by this lemma $\phi_G(X) = X - \{ \int_0^{X'} q(\theta) \, dG(\theta) \, dt / \int_0^{X'} q(\theta) \, dG(\theta) \},$

$$
|t_n(X) - \hat{\phi}_n(X)|^2 = \left| (\int_0^{X'} q(\theta) \, dG(\theta) \, dt / \int_0^{X'} q(\theta) \, dG(\theta)) \right|^2.
$$

Letting $\text{rhs}(1.5) = |(v/w) - (\hat{v}/\hat{w})|^2$ and $*$ denoting conditioning on $X$ and $(L(G,G_n) \leq \varepsilon)$ we apply Lemma 1.1.

$$
\text{rhs}(1.4) \leq 8E[|w|^2 E_{\hat{\phi}} |v - \hat{v}|^2] + 12E[|w|^2 E_{\hat{\phi}} |w - \hat{w}|^2].
$$

To get a bound for the second term of $\text{rhs}(1.4)$ we bound $\text{rhs}(1.6)$. To do so, we shall furnish the following two lemmas:

**Lemma 1.4.** For any $0 \leq \xi \leq 1/3$ and with $z=X'+t$ for $0 \leq t \leq 1$,

$$
\varepsilon[|G(z) - \hat{G}_n(z)|/w^2] \leq C(1+28)\varepsilon
$$

where $B$ and $C$ are positive constants in A i) and A ii).
Proof. By the definition of $L(G, \hat{G}_n)$ and by the fact that the infimum in the definition of Lévy distance is attained, we have

$$|\langle G - \hat{G}_n \rangle(z)| \leq \varepsilon + G(z+\varepsilon) - G(z-\varepsilon).$$

Thus,

$$\text{lhs}(1.7) \leq \varepsilon E(w^{-2}) + E[\{G(z+\varepsilon) - G(z-\varepsilon)\}/w^2].$$

But, since the second term of \(\text{rhs}(1.9) = \int \{G(x'+t+\varepsilon) - G(x'+t-\varepsilon)\} f(x)/w \, dx\), by the Taylor expansion applied to \(G(z+\varepsilon) - G(z-\varepsilon)\) and by two applications of \(f \leq 1\) for the first inequality and an application of Aii) for the second inequality,

$$\text{the second term of rhs}(1.9) = \int \left[ \frac{G'(y)}{w} \right] dy f(x)/w \, dx \leq 2\varepsilon B \int [G(x) - G(x')]^{-1} \, dx \leq 2\varepsilon BC.$$

On the other hand, since $E(w^{-2}) = \int w^{-1} f(x) \, dx$, by two applications of $f \leq 1$ an application of Aii) we have

$$E(w^{-2}) \leq C.$$

This and (1.10) gives us the asserted lemma.

Lemma 1.5. For any $0 < \varepsilon < 1/3$ and with $z = X' + t$ and $u = X' + s$ for $0 \leq t, s \leq 1$,

$$E_{\xi}^* \left[ \left| \frac{\langle G - \hat{G}_n \rangle(z) - \langle G - \hat{G}_n \rangle(u)\rangle}{w^2} \right| \right] \leq \varepsilon^2 C(1+2B)^2.$$

Proof.) Let $h_a^b = h(b) - h(a)$ until the end of the proof. As in the proof of Lemma 1.4, for $0 < \varepsilon < 1/3$,

$$\left| \langle G - \hat{G}_n \rangle(z) \right| \left| \langle G - \hat{G}_n \rangle(u) \right| [L(G, \hat{G}_n) \leq \varepsilon^2 + \varepsilon \{G \}_{ \varepsilon}^{u \varepsilon} + G \}_{ \varepsilon}^{u \varepsilon} + G \}_{ \varepsilon}^{z \varepsilon}]

Thus, in the similar fashion to (1.9) and by applying (1.10) and (1.11), we have

$$\text{lhs}(1.12) \leq \varepsilon^2 C + 2\varepsilon B \varepsilon + E[G \}_{ \varepsilon}^{u \varepsilon} + G \}_{ \varepsilon}^{z \varepsilon} /w^2\].$$

But, again, applying the similar method to obtaining (1.10) leads to

$$\text{(the third term of the extreme rhs}(1.14)) \leq 4\varepsilon^2 B^2 C$$

which completes the proof.

We shall now find an upper bound for both terms in the rhs(1.6).
Lemma 1.6. For $0 \leq t \leq 1$,

$$\mathbb{E}_n \{ (\int_{X'}^{X'+t} q(\theta) \, d\hat{G}_n(\theta) - \int_{X'}^{X'} q(\theta) \, dG(\theta))^2 / \omega^2 \} \leq m^2 \varepsilon^2 (m + 2(m+1)(1+2B))^2.$$  

Proof. Let the numerator in the curly bracket be $|Y-y|^2$. By letting $I=\{X', X'+t\}$, define by $G_I$ the retraction of $G$ into the closed interval $[G(X'), G(X)]$. Then, by Lemma 1.2, $L(G, I, \hat{G}_n) \leq L(G, \hat{G}_n) \leq \varepsilon_{SVT}$ where $S=|(G-\hat{G}_n)(X')|$ and $T=|(G-\hat{G}_n)(X'+t)|$. Thus,

$$L(G, \hat{G}_n) \leq \varepsilon \implies L(G, I, \hat{G}_I) \leq \varepsilon_{SVT} (\varepsilon \lambda).$$

By applying Lemma 1.3 with $h(\theta)$, the retraction of $q(\theta)$ to $I$, and weakening the resulted bound,

$$L(G, I, \hat{G}_I) \leq \lambda \implies |Y-y|^2 \leq (\alpha(\lambda+) + m(S + T))^2.$$  

To bound $\alpha(\lambda+)$, pick $\omega_1, \omega_2 \in I$ such that $0 < \omega_2 - \omega_1 < \lambda$. Now, by the definition of $h$ and $q$, $h(\omega_2) - h(\omega_1) = q(\omega_2) - q(\omega_1) = q(\omega_1) (\int_{\omega_1}^{\omega_2} f(s) \, ds - \int_{\omega_1}^{\omega_2+1} f(s) \, ds)$ and since $q \leq m$ and $f \leq 1$, $|h(\omega_2) - h(\omega_1)| \leq m^2 \lambda$. Thus, by Definition 1.1, $\alpha(\lambda) \leq m^2 \lambda$. Thus, the same bound applies for $\alpha(\lambda+)$. Therefore, by (1.17) and weakening the bound

$$L(G, I, \hat{G}_I) \leq \lambda \implies |Y-y|^2 \leq (m^2 \varepsilon + m(m+1)(S+T))^2.$$  

Thus, in view of (1.16) and (1.18),

$$\text{lhs(1.15)} \leq m^4 \varepsilon^2 E(w^{-2}) + 2m^3 (m+1) \varepsilon \{ E(S/w^2) + E(T/w^2) \} + m^2 (m+1)^2 \{ E(S^2/w^2) + 2E(ST/w^2) + E(T^2/w^2) \}.$$  

Thus, applying (1.11), Lemma 1.4 twice and Lemma 1.5 three times, leads to the asserted bound.
From (1.4), (1.6) and Lemma 1.6 we obtain

Theorem 1.1. If \( P_\theta \leq P(f) \) for \( (0, \infty)^{-1} \leq f \leq 1 \) and \( G \) satisfies the assumptions Ai) and Aii), then for \( 0 < \epsilon < 1/3 \),

\[
R(\hat{\theta}_n, G) - R < E[\log(G, \hat{G}_n) > \epsilon] + c_0 \epsilon^2
\]

where \( c_0 = 20m^2C(m+2(m+1)(1+2B))^2 \).

3. EB bootstrap estimates \( \hat{\theta}_n \) and \( \hat{\phi}_n \).

We first construct normalized (but not monotonized) estimates \( \hat{G} \) and \( G^*_n \) of an unknown prior \( G \) and then get a distribution-valued (monotonized) estimates \( \hat{G}_i \) and \( \hat{G}_n \) for \( G \). We then play Bayes vs \( \hat{G}_i \) and \( \hat{G}_n \) to get \( \hat{\theta}_n \) and \( \hat{\phi}_n \), respectively.

Let \( \Omega = (\infty, \infty) \). Let \( Q \) be the distribution function defined by

\[
Q(y) = \int_{-\infty}^{y} q(\theta) \, dG(\theta) \quad \text{for every } y.
\]

Since \( q \geq 1 \) and \( q \) is the density of \( Q \) wrt \( G \), it follows by Theorem 32.B of Halmos (1950) that

\[
G(y) = \int_{-\infty}^{y} q(\theta) \, dG(\theta).
\]

This form gives an estimate of \( G \) by inserting an estimate for \( Q \) in place of \( Q \).

Letting \( p(y) = \int p_\theta(y) \, dG(\theta) \), we have by the definition of \( p_\theta \) that \( p(y) = f(y)(Q(y) - Q(y')) \) and by a telescopic series,

\[
Q(y) = \sum_{r=0}^{\infty} \frac{p(y-r)}{f(y-r)}.
\]

We remark that if the \( r \)-th term of rhs(3.2) is nonzero, then

\[
r-1 \leq \text{range of } \{\theta_1, \theta_2, \ldots, \theta_n\}.
\]

We shall first exhibit \( \hat{\theta}_n \). To estimate \( Q \), we introduce a kernel function which is a nonnegative bounded-variation function vanishing outside the interval \( (0,1) \) and satisfying
(3.4) \[ 0 \leq K(y) \leq 1 \quad \text{and} \quad \int_0^1 K(y) \, dy = 1. \]

Let us first estimate \( p(y) \) by \( \hat{p}(y)=\left(\frac{nh}{n}\right)^{-\frac{1}{2}} \sum_{j=1}^n K\left(\frac{1}{n}X_j-y\right) \) where \( 0<h=h_n \to 0 \) as \( n \to \infty \). In view of (3.2), \( Q(y) \) is estimated by

(3.5) \[ \tilde{Q}(y) = \sum_{r=0}^{\infty} \frac{\hat{p}(y-r)}{f(y-r)}. \]

Thus, by (3.1) we estimate \( G(y) \) by

(3.6) \[ \tilde{w}(y) = \int_{-\infty}^{y} (q(\theta))^{-1} \, d\theta. \]

We use a raw estimate of the empirical distribution \( G_n \) of \( n \) parameters \( \theta_1, \ldots, \theta_n \) from \( n \) populations as an estimate of \( G \). Let \( F_n(y) = \frac{1}{n} \sum_{j=1}^n [X_j \leq y] \). Since \( F_n(y) \leq G_n(y) \leq F_n(y+1) \), we get a raw estimate

(3.7) \[ \hat{G}_n(y) = (F_n(y)\tilde{w}(y))^{-1} F_n(y+1). \]

We let \( \delta = N^{-1} \), \( N \) being a positive integer depending on \( n \), and consider the following grid points on the real line: \( \ldots < -2\delta < -\delta < 0 < \delta < 2\delta < \ldots \). We finally estimate \( G \) at \( y \) by

(3.8) \[ \hat{G}_n(y) = \sup \{ \hat{G}_n(j\delta) : j\delta \leq y \}, \]

Our EB bootstrap estimate \( \hat{\Phi}_n(X) \) at \( X \) (for \( \theta \)) is of form

(3.9) \[ \hat{\Phi}_n(X) = \int_{X_1}^X q(\theta) \, d\hat{G}_n(\theta) / \int_{X_1}^X q(\theta) \, d\hat{G}_n(\theta). \]

Similarly, we shall exhibit \( \hat{\Phi}_n \). To estimate \( Q \), we use above kernel function without the assumption \( \int_0^1 K(y) \, dy = 1 \) in (3.4). In view of (3.2) \( Q(y) \) is estimated by

(3.10) \[ Q^*(y) = \sum_{r=0}^{\infty} \left(\frac{nh}{n}\right)^{-\frac{1}{2}} \sum_{j=1}^n K\left(\frac{1}{n}X_j-y+r\right) / f(X_j) \}

Thus, by (3.1) we estimate \( G(y) \) by

(3.11) \[ \tilde{w}^*(y) = \int_{-\infty}^{y} (q(\theta))^{-1} \, dQ^*(\theta). \]
Hence, as above we get a raw estimate

\[(3.12) \quad G^*(y) = (F_n(y) \overline{W}(y))_{i \leq n}(y+1)\]

and finally estimate \( G \) at \( y \) by

\[(3.13) \quad \hat{G}_n(y) = \sup_G\{G^*(j\delta): j \leq y, j=0, \pm 1, \ldots\}.\]

The second EB estimate (for \( \theta \)) \( \hat{\theta}_n(X) \) at \( X \) is of form

\[(3.14) \quad \hat{\theta}_n(X) = \int_X \theta \phi(\theta) \, d\hat{G}_n(\theta)/\int_X \phi(\theta) \, d\hat{G}_n(\theta).\]

4. A Rate \( O(n^{-1/2}\log n)^{1/2} \) for \( R(\hat{\phi}_n, G)^{-1} \).

To get a rate of convergence for \( R(\hat{\phi}_n, G)^{-1} \) we use the bound of Theorem 1.1 we shall get an upper bound for \( \mathbb{E}[L(G, \hat{G}_n)^{-1}] \) (forthcoming Lemma 4.6). This bound is essentially given by bounding \( 1-\mathbb{E}[\{G(y-e) \leq \hat{G}_n(y) \leq G(y+e)\}] \) (forthcoming Lemma 4.5). Therefore, main part in this section is Lemma 4.5. But, since we will apply Theorem 2 of Hoeffding (1963) for its proof, we need to get the bounds for \( \mathbb{E}\hat{W}(y)\); with a positive constant \( b_{1,} \), \( G(y+h)+b_{1,} h \) for an upper bound and \( G(y)-b_{1,} h \) for a lower bound (Lemma 4.4). To do so we shall furnish Lemmas 4.1, 4.2 and 4.3.

Throughout this section we assume \( \Omega = [c, d] \) where \(-\infty < c < d < \infty\). In addition to the assumption on \( f \) in Section 2 we also assume that \( 1/f \) satisfies the Lipshitz condition:

\[(4.1) \quad \sup_{|v-u|} \|f(v)^{-1} - f(u)^{-1}\|_{u<v} \leq M\]

for a finite nonnegative constant \( M \). This is equivalent to

\[(4.2) \quad |(f(s)/f(t)) - 1| \leq M|s-t|.

By the definition (3.7) of \( \overline{W}, \overline{W} = \sum_{j=1}^{n} \overline{W}_j, \) where for each \( j \)
where the subscript \( t \) in \( d_t \) denotes the variable of integration. To find bounds for \( \mathbb{E}(y) \) we shall find an upper and a lower bound for \( \tilde{W}_j \). Fix \( j \) and use the corresponding notations without subscript \( j \) until the end of the proof of Lemma 4.4. Hereafter, we abbreviate \( g(b) - g(a) \) to \( g_a^b \) until the end of the proof of Lemma 4.5.

**Lemma 4.1.** For \( 0 < h < 1 \),

\[
(4.4) \quad |h(q(\theta))^{-1}P_{\theta}\tilde{W} - I(y) + S(y)| \leq Mh^2(1+(M\lambda)N)
\]

where 

\[
N = d+2, \quad I(y) = (h(q(\theta)))^{1/0} K(u)[\theta-y\in hu] \, du \quad \text{and} \quad S(y) = h^{1/0} \int_0^1 K(v)[\theta-t\in hv] \, dv
\]

**Proof.** Because a function satisfying the Lipschitz condition is absolutely continuous (cf. Ryden (1968), p.108)) and \( 1/q \) is clearly absolutely continuous, \( 1/f(-r) \), \( 1/q \) and \( K \) are all of bounded variation in the definition (4.3) of \( \tilde{W} \).

Applying integration by parts (Saks (1937), Theorem III. 14. 1) and using \( d(q(t))^{-1} = f_t^{1+1} \) gives us that

\[
(4.5) \quad \tilde{W} = \sum_{r=0}^\infty \frac{K(h^{-1}(x+r))}{f(y-r)q(y)} - \int_{-\infty}^y \sum_{r=0}^\infty \frac{K(h^{-1}(x+r))}{f(t-r)} f_t^{1+1} \, dt.
\]

Taking expectation wrt \( P_{\theta} \) and multiplying \( (q(\theta))^{-1} \) on both sides gives

\[
(4.6) \quad h(q(\theta))^{-1} P_{\theta} \tilde{W} = (q(y))^{-1} \sum_{r=0}^\infty \int_{y-r}^{y-r+h} [\theta-s<\theta+1] K(h^{-1}(s-y+r)) \frac{f(s)}{f(y-r)} \, ds
\]

\[
- \int_{-\infty}^y \sum_{r=0}^\infty \int_{t-r}^{t-r+h} [\theta-s<\theta+1] K(h^{-1}(s-t+r)) \frac{f(s)}{f(t-r)} \, ds \, f_t^{1+1} \, dt.
\]

Changing the variables \( u = h^{-1}(s-y+r) \) and \( v = h^{-1}(s-t+r) \) for the integrals in the first term and the second term, respectively, we get
\[(4.7) \quad \text{rhs}(4.6) = h(q(y)\int_{0}^{\infty} K(u)[\theta-(y-r)\leq hu<\theta+1-(y-r)] \frac{f(y-r+hu)}{f(y-r)} \, du
\]
\[- h^{\prime} \int_{0}^{\infty} K(v)[\theta-(t-r)\leq hv<\theta+1-(t-r)] \frac{f(t-r+hv)}{f(t-r)} \, dv \int_{0}^{t+1} dt\]

Since by (4.2) \(|f(t-r+hv)/f(t-r)| \leq \text{Mh} \) for \(0 \leq v \leq 1\), applying telescopic series gives
\[(4.8) \quad \left|h(q(\theta))^{-1} P_{\theta} \tilde{W} - I(y) + S(y)\right| \leq \text{Mh}^2 \int_{0}^{1} K(u)[\theta-y<hu] \, du
\]
\[+ \int_{0}^{1} K(v)[\theta-t<hv] \, dv \int_{0}^{t+1} dt\].

Since \((0 \leq K \leq 1)\), \([\theta-t<hv] \leq [\theta-t<\theta]\), \(|f(t+1) \leq \text{M} \lambda^2\) and \((q(y))^{-1} \leq 1\), \(\text{rhs}(4.8) \leq \text{Mh}^2 (1+\left(\text{M} \lambda\right)(y+\theta+h))\) which proves the lemma.

Following Lemma 4.2 is straightforward, so we omit the proof.

**Lemma 4.2.** For \(I(y)\) defined in Lemma 4.1,
\[(4.9) \quad h(q(\theta))^{-1} [\theta-y] \leq I(y) \leq h(q(\theta))^{-1} [\theta+y+h].\]

**Lemma 4.3.** For \(S(y)\) defined in Lemma 4.1
\[(4.10) \quad |S(y) - h[\theta+y](q(y))^{-1} - (q(\theta))^{-1}| \leq (\text{M} \lambda) h^2.\]

**Proof.** Since by (3.4) \(\int_{0}^{\infty} K(u) \, du = 1\),
\[(4.11) \quad S(y) = h[\int_{0}^{\infty} [\theta-t] f(t+1) \, dt + \int_{0}^{\infty} [\theta-h-t<\theta] f(t+1) \, dt].
\]
Since \(\int_{0}^{\infty} [\theta-t] \, f(t+1) \, dt = [\theta-y] \left[1/q(y) - 1/q(\theta)\right]\), using \(|f(t+1) \leq \text{M} \lambda^2\) and \((0 \leq K(t) \leq 1)\),

and weakening the resulting bound leads to the asserted bound.

We are now ready to prove the following lemma:

**Lemma 4.4.** For any \(y \in [c-1, d+1]\),
\[(4.12) \quad G(y) \leq b_1 h \leq M(y) \leq G(y+h) + b_1 h\]
where \(b_1 = (\text{M} \lambda + \text{M}(1+\text{M} \lambda)) + h \lambda ((1+h) \lambda m) \).

**Proof.** From Lemmas 4.1, 4.2 and 4.3 and after a little computation,
\[(4.13) \quad [\theta-y]/q(\theta) - b_2 h \leq P_{\theta} \tilde{W}/q(\theta) \leq [\theta-y+h]/q(\theta) + [y \leq \theta+y+h] (q(y))^{-1} (q(\theta))^{-1} + b_2.\]
where \( b_2 = \{(M \mid 1) + M(1+N(M \mid 1))\} \). From the definition and \( f \leq 1 \),

\[
(4.14) \quad [y \leq \theta \leq y+h] |(q(y))^{-1} - (q(\theta))^{-1}| \leq h.
\]

On the other hand, \( \text{lhs}(4.14) = [y \leq \theta \leq y+h] \left\{ (q(\theta)/q(y))^{-1} - 1 \right\} \leq ((1+h)M \lambda m \right\} \leq ((1+h)M \lambda m \right\) \leq (\text{lhs}(4.14) \leq h \lambda ((1+h)M \lambda m)) \). Therefore, noticing \( q \leq m \) leads to

\[
[\theta \leq y] - b_2 h \leq E_\theta \tilde{W} \leq [\theta \leq y+h] + b_1 h
\]

where \( E_\theta \) means the expectation wrt \( n \) products of the conditional probability distribution \( P_\theta \) and \( b_1 = b_2 + 1/(1 + h) \lambda m \). Taking the expectation wrt \( G \) gives the asserted bound.

Following Lemma 4.5 becomes a direct generalization of Lemma 3.1 of Fox (1970) in the empirical problem in the sense that if \( f = 1 \), then \( m = 1 \) and \( M = 0 \) and hence we get his bound \( 2 \exp(-2nh^2) \).

**Lemma 4.5.** If \( 0 < h \leq 1 \), then for each \( y \)

\[
(4.15) \quad 1 - E(\{G(y-\epsilon) - \epsilon \leq \tilde{W}(y) \leq G(y+\epsilon) + \epsilon\}) \leq 2 \exp\left\{ - \frac{2nh^2((\epsilon - b_1 h)^2)}{(b_3 + 2b_4)^2} \right\}
\]

where \( b_3 = (1 + (N+1)M \lambda m (N+1)) \) and \( b_4 = (m(M \mid 1)) \lambda ((N+1)(1+MN/2)) \).

**Proof.** For any \( y \), it is sufficient to prove the lemma for the raw estimate \( \tilde{W} \), for if \( G(y-\epsilon) - \epsilon \leq \tilde{W}(y) \leq G(y+\epsilon) + \epsilon \), it is easily checked that \( G(y-\epsilon) - \epsilon \leq \tilde{W}(y) \leq G(y+\epsilon) + \epsilon \).

As in the proof of Lemma 3.1 of Fox (1970) we shall apply Theorem 2 of Hoeffding (1963). To do so we shall use the bounds for \( E_{\tilde{W}}(y) \) in Lemma 4.4 and furthermore need to get an upper and a lower bound of \( \tilde{W} \) for each \( j \).

In the definition (4.5) of \( \tilde{W} \) since \( K(h^{-1}(X-y+r)) \geq 0 \) when \( X-hsY-r \leq X \), there is at most one positive term in the first term of \( \text{rhs}(4.5) \). Applying (4.2) and the
fact that \( r \leq N \) and \( K \leq 1 \) gives that

\[
0 \leq \text{the first term of } \text{rhs}(4.5) \leq 1 + (N+1)M.
\]

Since \( f^{-1} \leq m, q^{-1} \leq 1 \) and \( K \leq 1 \), we also have the first term of \( \text{rhs}(4.5) \leq m(N+1) \).

Hence, putting these two bounds together gives

\[
(4.16) \quad 0 \leq \text{the first term of } \text{rhs}(4.5) \leq (1 + (N+1)M) \Lambda m(N+1).
\]

In the same way,

\[
(4.17) \quad \text{the second term of the } \text{rhs}(4.5)
\]

\[
= \int_{-\infty}^{\infty} \sum_{r=0}^{\infty} K_h^{-1}(X-t+r) \frac{f(t-r)}{f(t)} \text{dt}.
\]

Since the summation in \( r \) has at most one positive term and since \( |f(t)| \leq M/1 \), \( f^{-1} \leq m \) and \( K \leq 1 \), we have

\[
(4.18) \quad |(4.17)| \leq hM(N+1).
\]

On the other hand, by changing a variable \( u = h^{-1}(X-t+r) \)

\[
(4.19) \quad (4.17) = h \sum_{r=0}^{\infty} \int_{-\infty}^{\infty} h^{-1}(X+y+r) \left[ 0 < u \leq 1 \right] \frac{K(u)}{f(X-hu)} \left( f(X+r-hu+1) - f(X+r-hu) \right) du
\]

Since \( f(X+r-hu+1)/f(X-hu) \leq 1 + (r+1)M \) and \( K \leq 1 \), extending the range of integration leads to

\[
(4.20) \quad |(4.17)| \leq h(N+1)(1+MN/2).
\]

Therefore, putting two bounds (4.18) and (4.20) together and using \( h \leq 1 \) gives

\[
|\text{rhs}(4.17)| \leq b_4.
\]

Therefore, in view of (4.5), (4.16) and (4.17), \( -b_4 \leq h\tilde{W} \leq b_3 + b_4 \).

We now apply Theorem 2 of Hoeffding (1963). Since \( h \leq \varepsilon \), using the second inequality of (4.12) in Lemma 4.4 and applying Theorem 2 of Hoeffding (1963) gives

\[
(4.21) \quad \mathbb{E}[\tilde{W}(y) > G(y+\varepsilon) + \varepsilon] \leq \mathbb{E}[\tilde{W}(y) - \tilde{W}(y) > \varepsilon - b_1 h]
\]

\[
\leq \exp\left( -\frac{2nh^2((c-b_1 h)^+)^2}{(b_3 + 2b_4)^2} \right).
\]
Furthermore, by the first inequality of (4.12) \( \bar{W}(y) \leq G(y) - \varepsilon \), hence, by the symmetry of the tail bounds, \( \mathbb{E}[\bar{W}(y) < G(y) - \varepsilon] \leq \text{rhs}(4.21) \), which together with (4.21) gives us the asserted bound of Lemma 4.5.

With \( \hat{G}_n \) defined by (3.8) we can use Lemma 4.6 (Nogami (1982, Lemma 2.6)) below for an upper bound of the first term of the right hand side of Theorem 1.1. By the triangular inequality applied to the first inclusion sign,
\[
\{L(G, \hat{G}_n) > 2\varepsilon\} \subseteq \{L(G, \hat{G}_n) + L(G_n, \hat{G}_n) > 2\varepsilon\} \subseteq \{L(G, \hat{G}_n) > \varepsilon\} u \{L(G_n, \hat{G}_n) > \varepsilon\}
\]
where \( G_n \) is the empirical distribution of \( \theta_1, \ldots, \theta_n \). Hence,
\[
\mathbb{E}[L(G, \hat{G}_n) > 2\varepsilon] \leq \mathbb{E}[L(G, \hat{G}_n) + L(G_n, \hat{G}_n) > 2\varepsilon] \leq \mathbb{E}[L(G, \hat{G}_n) > \varepsilon] + \mathbb{E}[L(G_n, \hat{G}_n) > \varepsilon].
\]

We shall use Lemma 4.6 below to get a bound for the second term of the extreme \text{rhs}(4.22). Since the proof involves only on the truncation of the raw estimate \( \bar{W}(y) \) into \( F_n(y), F_n(y+1) \) and the monotonicity of \( \hat{G}_n \), the same proof as Lemma 2.6 of Nogami (1982) applies. Hence, we omit it.

**Lemma 4.6.** (Nogami (1982, Lemma 2.6)) For any \( \varepsilon > 0 \), if \( h \leq \varepsilon \) and \( \delta \leq \varepsilon \), then
\[
\mathbb{E}[L(G, \hat{G}_n) > 2\varepsilon] \leq (\delta^{-1} + 1)\varepsilon^{-1} + 1)(\text{rhs}(4.15)).
\]

For bounding the first term of the extreme \text{rhs}(4.22), we use Lemma 4.7 below.

**Lemma 4.7.**
\[
\mathbb{E}[L(G, G_n) > \varepsilon] \leq 2\exp(-2\varepsilon^2).
\]

**Proof.** Since the lhs(4.24) = \( \mathbb{E}[G(y) > G_n(y) + \varepsilon] + \mathbb{E}[G(y) < G_n(y) - \varepsilon] \) \leq \( \mathbb{E}[G(y) > G_n(y) + \varepsilon] + \mathbb{E}[G(y) < G_n(y) - \varepsilon] \leq \mathbb{E}[G(y) - G_n(y) > \varepsilon] + \mathbb{E}[G_n(y) - G(y) > \varepsilon]. \) Since \( \mathbb{E}[G_n(y)] = G(y) \), Hoeffding's inequality (1963) applied twice gives us the asserted bound of the lemma.
Lemmas 4.6 and 4.7 gives

\[ (4.25) \quad \mathbb{E}[\mathcal{L}(G, \hat{\phi}_n) > 2\varepsilon] \leq (\varepsilon^{-1} + 1)(\varepsilon^{-1} + 1)(\text{rhs}(4.15)) + 2\exp\{-2n\varepsilon^2\}. \]

To get a rate of convergence for \( R(\hat{\phi}_n, G) - R \) defined in (1.2) we use the bound of Theorem 1.1. Let us assume \( \Omega = [c, d] \) in the definition of \( p(f) \). Since the similar proof to that for Theorem 2.1 of Nogami (1982) does work, we shall omit the proof of the following theorem:

**Theorem 4.1.** If \( p_0 \geq p(f) \) with \( \Omega = [c, d] \) and \( f \) satisfying the assumptions 
\( (0 < m^{-1} \leq f \leq 1 \) and the Lipshitz condition \( (4.1) \) and if \( G \) satisfies the assumption \( \text{Ai} \) and \( \text{Aii} \), then there exist positive constants \( a_1 \) and \( a_2 \) so that, for \( \hat{\phi}_n \) with \( a_1 h = a_2 \hat{\phi} = (n^{-1} \log n)^{\frac{1}{2}} \),

\[ (4.23) \quad |R(\hat{\phi}_n, G) - R| = O\left( (n^{-1} \log n)^{\frac{1}{2}} \right). \]

5. **A Rate \( O\left( (n^{-1} \log n)^{\frac{1}{2}} \right) \) for \( R(\hat{\phi}_n, G) - R \).**

To get a convergence rate for \( R(\hat{\phi}_n, G) - R \) we shall proceed in the similar fashion to Section 4. In this section we dare not assume Lipshitz condition (4.1) or equivalently (4.2). As we can see from Section 4, Lipshitz condition (4.1) leads Lemma 4.5 to a direct generalization of Lemma 3.1 of Fox (1970). Even in this section we may be able to use this assumption (4.1) for this purpose. However, we may not have much interest of it, so that we do not assume (4.1).

By the definition (3.11) of \( \overline{W}_n, \overline{w}_n \), \( \overline{w}_n = n^{-1} \sum_{j=1}^{n} w_j \) where for each \( j \)

\[ (5.1) \quad h \overline{w}_j = \sum_{t=0}^{\infty} J(t)_{\overline{Y}} - (q(t))^{-1} \int F^{-1}(X_{t+} - t + r) \cdot d_t \{ K(h^{-1}(X_{t+} - t + r)/f(X_{t+})) \}, \]

where the subscript \( t \) in \( d_t \) is defined in (4.3). Fix \( j \) and use the corresponding notations without subscript \( j \) until the end of the proof of Lemma 5.3. Following Lemma 5.1 corresponds to Lemma 4.1.
Lemma 5.1. For \( 0 < h < 1 \),

\[
(5.2) \quad h(q(\theta))^{-1} p_{\theta}^b W^* = I(y) - S(y)
\]

where \( I(y) \) and \( S(y) \) are defined in Lemma 4.1.

**Proof.** Let \( g|_{a}^{b} = g(b) - g(a) \) until the end of the proof. By the similar reason to getting (4.5),

\[
(5.3) \quad h W^* = \sum_{r=0}^{\infty} \frac{K(h^{-1} (X-y+r))}{f(X)} q(y) f_{t}^{t+1} dt.
\]

Thus,

\[
(5.4) \quad h(q(\theta))^{-1} p_{\theta}^b W^* = (q(y))^{-1} \sum_{r=0}^{\infty} \int_{y-r}^{y-r+h} k(h^{-1} (s-y+r)) [0 \leq s < \theta + 1] ds
\]

\[
- \int_{y-r}^{y} k(h^{-1} (s-t+r)) [0 \leq s < \theta + 1] ds f_{t}^{t+1} dt.
\]

Changing the variables \( u = h^{-1} (s-y+r) \) and \( v = h^{-1} (s-t+r) \) in the first and second integration in the rhs gives

\[
(5.5) \quad \text{rhs}(5.4) = h(q(y))^{-1} \int_{0}^{1} K(u) [0 \leq y+ru < \theta + 1 - y+r] du
\]

\[
- \int_{0}^{1} K(v) [0 \leq v < \theta + 1 - t+r] dv f_{t}^{t+1} dt.
\]

Two telescopic series leads to the equality in the lemma.

We use Lemmas 4.2 and 4.3 to get following Lemma 5.2 which corresponds to Lemma 4.4.

**Lemma 5.2.** For any \( y \)

\[
G(y) - mh \leq E_{\theta} W^*(y) \leq G(y+h) + 2mh.
\]

**Proof.** By Lemmas 5.1, 4.2 and 4.3 with \( MA1 \) replaced by 1,

\[
[\theta \leq y] - h(q(\theta)) \leq P_{\theta}^b W^*
\]

\[
\leq [\theta \leq y+h] + q(\theta) [y \leq s \leq y+h] (q(y)^{-1} - (q(\theta))^{-1}) + q(\theta) h
\]

Thus, applying inequalities (4.14) and \( 0 < q \leq m \) and taking average wrt \( j \) and then expectation wrt \( \mathbb{E} \) leads to the bounds of the asserted lemma.
Following Lemma 5.3 is no more a direct generalization of Lemma 3.1 of Fox (1970). However, as the author stated at the beginning of this section, we can easily check that if we have the Lipshitz assumption (4.1) on \(1/f\), then resulted Lemma 5.3 would be a direct generalization of Fox's lemma.

**Lemma 5.3.** If \(0 < h \leq 1\), then for every \(y\)

\[
(5.6) \quad 1 - \mathbb{P}(G(y - \epsilon) - \epsilon \leq G^*(y) \leq G(y + \epsilon) + \epsilon) \\
\leq 2\exp\left(\frac{-2nh^2((c-2mh)^+)^2}{(1+2h)m^2}\right)
\]

**Proof.** The proof is similar to that for Lemma 4.5 except bounds for \(W^*_j\).

In view of (5.1), we have that \(0 \leq\) the first term of \(\text{rhs}(5.1) \leq m\) because \((0 \leq) k \leq 1, f^{-1} \leq m\) and \(q^{-1} \leq 1\). Similarly, (4.18) with \(M\|1\) replaced by 1, gives us the second term of \(\text{rhs}(5.1) \leq h m\). Therefore,

\[-m \leq W^* \leq (h^{-1} + 1)m.\]

Using Lemma 5.2 and Hoeffding's bound (1963) gives the asserted bounds.

As in (4.25), Lemmas 5.3 and 4.7 gives

\[
(5.7) \quad \mathbb{P}(L(G, \hat{G}_n) > 2\epsilon) \leq (\delta^{-1} + 1)(\epsilon^{-1} + 1)(\text{rhs}(4.6)) + 2\exp\left(-2n\epsilon^2\right).
\]

Therefore, we obtain without a proof

**Theorem 5.1.** If \(p \in p(f)\) with \((0 <) m^{-1} \leq f \leq 1\) and if \(G\) satisfies A1) and A1i), then there exist positive constants \(k_1\) and \(k_2\) so that for \(k_1 h = k_2 \delta = (n^{-1} \log n)^{\frac{1}{2}}\),

\[
(5.8) \quad |\hat{R}_{n}(G) - R| = O((n^{-1} \log n)^{\frac{1}{2}}).
\]
REFERENCES


