# Differences in the genetic structure between and within two landlocked Ayu groups with different migration patterns in Lake Biwa revealed by environmental DNA analysis 

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#### Abstract

Ayu (Plecoglossus altivelis) is largely an annual amphidromous fish, although a landlocked population lives in Lake Biwa, the largest lake in Japan. The landlocked population comprises two migrant groups, spring migrants and autumn migrants, which run to inlet rivers from the lake at different seasons. We used environmental DNA (eDNA) analysis, which is reported to be more sensitive and cost-effective than capture surveys, to clarify the genetic structure of this landlocked Ayu population with different migration patterns in Lake Biwa. We took water samples in 11 inlet rivers in the spring and autumn for 2 years in a row and quantitatively detected a total of 265 haplotypes of the mitochondrial D-loop region. The pairwise fixation index ( $F_{\mathrm{ST}}$ ) value and haplotype diversity indicated that there were genetic differences between the two migrant groups in their respective rivers, and the $F_{\mathrm{ST}}$ values were negatively related to latitude and the presence of artificial fish stocking. Additionally, isolation by distance within spring migrant group was observed when the lake was divided into the east and west sides. These findings show that the landlocked Ayu population in Lake Biwa has genetic structure associated with migration patterns and geographical distance. This study demonstrates that the eDNA approach will be effective for conducting a large-scale investigation of genetic structure beyond simple presence/absence tests.


## KEYWORDS

environmental DNA, genetic structure, Lake Biwa, landlocked Ayu, migration pattern

## 1 | INTRODUCTION

Population genetic studies of aquatic macroorganisms provide key information such as genetic structure, diversity, effective population sizes, and the extent of gene flow between populations for making strategies for their effective conservation and management (Bravington et al., 2016; Ruzzante et al., 2019; Székely et al., 2021). However, these studies generally require collection of an adequate number of tissue samples from multiple geographically distant sites
where the target species inhabits. This requirement may be potentially harmful to target species and their habitats when targeting rare or endangered species, and, in contrast, inadequate sampling may not provide accurate population genetic information when targeting species with large biomass and genetic diversity (Fosså et al., 2002; Hansen, 1988; Jørstad, 2004; Pirhonen \& Schreck, 2003; Xing et al., 2013).

Recently, environmental DNA (eDNA) analysis has attracted attention due to its great potential utility in assessing genetic

[^0]diversity beyond species detection (e.g., Andres et al., 2021; Baker et al., 2018; Marshall \& Stepien, 2019; Sigsgaard et al., 2016; Stat et al., 2017; Stepien et al., 2019; Yoshitake et al., 2019). For example, an eDNA-based mitochondrial DNA (mtDNA) survey using 1-L water samples provided comparable results from the conventional capture-based method with the Sanger sequencing for a large number (~100) of specimens (Tsuji, Maruyama, et al., 2020). In addition, an eDNA-based survey enables reasonable estimation of genetic indicators (e.g., nucleotide diversity, Tajima's test statistic, and effective female population size) by using the relative number of DNA copies of each mtDNA haplotype amplified from water samples (Tsuji, Shibata, et al., 2020). Thus, the use of eDNA analysis could provide the opportunity to collect population genetic data efficiently and noninvasively.

Ayu (Plecoglossus altivelis; Plecoglossidae, Teleostei) is a major member of stream communities in Japan occupying a unique ecological niche as an algivorous specialist and has long been commercially important as a fishery resource (Iguchi et al., 2002; Kawanabe, 1996; Miyadi, 1960). Mostly, Ayu has an amphidromous life history, in which the larvae inhabit the coastal zone and migrate to rivers from the sea for further growth, maturation, and reproduction throughout a 1-year life cycle (Nishida, 2001). On the other hand, landlocked populations of this species are found in a few lakes, including Lake Biwa, the largest lake in Japan. The population in Lake Biwa is enormous, several hundred to 2000 metric tons of this fish being annually harvested (Ministry of Agriculture, Forestry and Fisheries, 2021). Intriguingly, the majority of the population's life history is not amphidromous but anadromous-like (cf. Azuma, 1970, 1973a; Matsuyama \& Matsuura, 1985); a major part of fish enter inlet rives only for spawning.

The different life-history patterns in the landlocked Ayu population in Lake Biwa have long been of interest in the fields of fisheries, ecology, and evolution (Kawanabe, 1985, 1996). The population is typically divided into two migrant groups based on the run timing to inlet rivers from the lake (Figure 1); (1) the spring migrants (late MarchApril; so-called O-Ayu), which grow in the middle to upper reaches of the inlets and spawn at the lower reaches (i.e., amphidromouslike) and (2) the autumn migrants (late August-September; socalled Ko-Ayu), which spend nearly their entire lives in the lake and migrate to the inlets to spawn in autumn (i.e., anadromouslike) (Azuma, 1973b; Iguchi et al., 2002; Iguchi \& Nishida, 2000). With the exception that the nursery ground is a freshwater lake rather than the sea, the former's life history is essentially identical to that of other river populations of this species. Spring migrants feed on epilithic algae on the river bed (vs. zooplankton in the lake; Azuma, 1973a, 1973b) and finally grow much larger than the autumn migrants (standard length $\pm$ SD: spring migrants, $17.1 \pm 0.7 \mathrm{~cm}$; autumn migrants, $9.2 \pm 0.4 \mathrm{~cm}$; Iguchi, 1996). On the other hand, the autumn migrants have much larger population size than the spring migrants (cf. Azuma, 1970, 1973a, 1973b, 1973c). Furthermore, the autumn migrants tend to spawn earlier (late August-late September) than the spring migrants (middle September-middle November) (Azuma, 1973b; Tsukamoto \& Kajihara, 1987).

(b)
(1) Are there genetic differences between the two migrant groups?

(2) Are there differences in the spatial and temporal genetic structure between and within the migrant groups?


FIGURE 1 (a) Differences in the migration timing and growth areas in the two migrant groups (spring and autumn migrants) of the Ayu population landlocked in Lake Biwa and (b) the two hypotheses to be tested in this study.

Diversification of migration patterns often facilitates reproductive isolation between migrant groups, which may lead to population differentiation (Hendry et al., 2000; Moser et al., 2012; Ricker, 1940). This could be due to several factors: (1) the difference in body size between different types could drive a sexual assortative mating through mate choice (Foote \& Larkin, 1988; Moser et al., 2012); (2) the shift in reproductive timing associated with the divergence could cause phenological assortative mating (e.g., Feder et al., 1997; Santos et al., 2011). These are both plausible in the case of Ayu in Lake Biwa; it tends to choose mating partners of similar size as themselves (Iguchi \& Maekawa, 1993; Nishida et al., 1974),
and autumn migrants tend to spawn earlier than spring migrants (Azuma, 1973b; Tsukamoto et al., 1987). These facts lead to the idea that some extent of reproductive isolation has developed between the two migrant groups and they might be in the early phases of population differentiation. If it is the case, the landlocked Ayu in Lake Biwa will provide an excellent opportunity to observe the process of population differentiation that occurs as a result of landlocking.

The only prior study that investigated genetic variation between and within the two migrant groups found that there were tiny but significant genetic differences between the two groups ( $F_{\text {ST }}<0.34$ in the mtDNA D-loop region; Iguchi et al., 2002). Also, within the autumn migrants, there was a tendency toward isolation by distance. However, because of the limited sample size in both the specimens and rivers analyzed, further study with more sophisticated sampling designs both spatially and temporally is needed (Iguchi et al., 2002). If the landlocked Ayu in Lake Biwa is not genetically homogeneous as suggested by Iguchi et al. (2002), it should be considered in fisheries management including fishing, stocking, and seed production in aquaculture. For example, although fish stocking is an effective way to reduce anthropogenic impacts on critical fishery resources, it could also impose ecological and genetic risks if genetically improper fish are stocked (Iguchi, 2011; Laikre et al., 2010). In Lake Biwa, a large number of juveniles captured from the lake are stocked in the middle to upper reaches of several inlet rivers every year in early May without consideration of genetic aspects (Azuma, 1973b; Iguchi et al., 2002). Currently, the magnitude of impacts of stocking on population dynamics and genetic structure of the population cannot be accurately examined due to a lack of information about the exact genetic structure of landlocked Ayu population in Lake Biwa.

The purpose of this study was to clarify the genetic structure of the landlocked Ayu population in Lake Biwa with different migration patterns. To perform a rapid and geographically wide-ranging survey, we adopted eDNA for the evaluation of genetic diversity and differentiation. To achieve the goal of this study, we test the following two questions using population genetic parameters estimated by eDNA analysis: (1) whether genetic differences exist between the two migrant groups, as shown by limited sampling and data in Iguchi et al. (2002) and (2) whether there are differences in the spatial and temporal genetic structure between the migrant groups. Based on results, we explored the factors of genetic differentiation between migrant groups and geographical population structure, with particular attention to spawning timing and lake currents.

## 2 | MATERIALS AND METHODS

## 2.1 | Study sites

Genetic structure of the two migrant groups of the landlocked Ayu was examined using an eDNA-based method in 11 inlet rivers of Lake Biwa in Shiga Prefecture, central Honshu, Japan. Water sampling was conducted in the lower reaches of 11 rivers (Figure 2); Echi River (1-EC), the Uso River (2-US), the Inukami River (3-IN), the Seri


FIGURE 2 The studied rivers and gyre system in Lake Biwa. The solid and open circles indicate the rivers where fish have been stocked or not, respectively.

River (4-SR), the Amano River (5-AM), the Shiotsu-o River (6-SO), the Chinai River ( $7-\mathrm{CN}$ ), the Ishida River (8-IS), the Ado River (9-AD), the Wadauchi River (10-WD), and the Wani River (11-WN). A large number of juveniles captured from Lake Biwa were stocked into the middle and/or upper reaches in 1-EC, $3-\mathrm{IN}, 8$-IS and 9-AD every spring. Until 2008, stocking was done in 4-SR.

## 2.2 | Water sample processing and eDNA extraction

Water sampling was performed in each river in spring and autumn in 2018 and 2019 (April 22 and September 14, 2018 and April 19 and August 27, 2019). By collecting water samples from the lower reaches of rivers at the time when each migrant group migrates (i.e., Spring or Autumn), we can collect eDNA derived from individuals belonging to each group. Autumn water sampling was performed before the spring migrants moved from the upper to lower reaches of the river to spawn. A total of about $4 L$ of surface water was collected from three to five points in the transverse direction at each sampling site using a disposable plastic cup. The collected water was pooled and carefully agitated in a bucket. The water qualities of each sample are listed in Table S1. Of about 4 L of water sample, 3 L was filtered using three Whatman GF/F filters
(1 L/filter, $0.7 \mu \mathrm{~m}$ pore size; GE Healthcare) on-site. To monitor contamination during the filtering and subsequent DNA extraction steps, 1 L of ultrapure water was filtered on-site in the same manner. The filter samples were placed in a plastic bag and immediately kept at $-20^{\circ} \mathrm{C}$. All filtering equipment was decontaminated by soaking in $10 \%$ bleach solution water before use. DNA was extracted from a total of 144 filter samples (132 real samples, 12 filter negative controls) using the spin column (EP-31201; GeneDesign) and DNeasy Blood and Tissue Kit (Qiagen). At the final extraction step, DNA was eluted from the column with $100 \mu \mathrm{l}$ of Buffer AE. Detailed information about DNA extraction is described in Appendix S1.

## 2.3 | Paired-end library preparation and sequencing with MiSeq

To quantitatively evaluate the concentration of each Ayu haplotype, we adopted the quantitative eDNA metabarcoding method using the internal standard DNAs (qMiSeq method, Ushio et al., 2018). The use of the qMiSeq method enables us to estimate eDNA copy numbers of detected unknown Ayu haplotypes by using sample-specific standard lines which were obtained for each sample based on the relationship between the number of sequence reads and the known copy number of standard DNAs (Table S2: Tsuji, Shibata, et al., 2020; Ushio et al., 2018). In the DNA library preparation, we employed a two-step tailed PCR approach to construct the paired-end libraries.

The first-round PCR (first PCR) was carried out with a $12-\mu \mathrm{l}$ reaction volume to amplify the Ayu mitochondrial D-loop region and the standard DNAs (insert length, 166 bp , both). The sequences of primers are: PaaDlp-2_F (5'-ACA CTC TTT CCC TAC ACG ACG CTC TTC CGA TCT NNN NNN CCG GTT GCA TAT ATG GAC CTA TTA C$3^{\prime}$ ), PaaDIp-2_R1 and R2 (5'-GTG ACT GGA GTT CAG ACG TGT GCT CTT CCG ATC TNN NNN NGC TAT TRT AGT CTG GTA ACG CAA G-3'). The second-round PCR (second PCR) was carried out with a $12-\mu$ I reaction volume to add Illumina sequencing adaptors and the dual-index (i.e. barcode) sequences. The sequences of primers are: Forward (5'-AAT GAT ACG GCG ACC ACC GAG ATC TAC AXX XXX XXX ACA CTC TTT CCC TAC ACG ACG CTC TTC CGATCT-3'), and Reverse ( $5^{\prime}$-CAA GCA GAA GAC GGC ATA CGA GAT XXX XXX XXG TGA CTG GAG TTC AGA CGT GTG CTC TTC CGA TCT-3').

All indexed products of the second PCR were pooled in equal volumes, and the target size of the libraries (ca. 370-bp) was excised using $2 \% \mathrm{E}-\mathrm{Gel}$ SizeSelect Agarose Gels (Thermo Fisher Scientific). The excised library was adjusted to 4 nM and sequenced on the MiSeq platform using MiSeq v2 Reagent Micro or Nano Kit for $2 \times 150$ bp PE cartridges (Illumina) with $30 \%$ PhiX spike-in according to the manufacturer's instructions. All raw sequences obtained in this study were deposited in the DDBJ Sequence Read Archive (accession number: DRA013835). Detailed information about pairedend library preparation and sequencing by MiSeq are described in Appendix S1.

## 2.4 | Bioinformatic analysis

To perform the denoising of erroneous sequences, the Fastq files containing raw reads were processed using the Divisive Amplicon Denoising Algorithm 2 package ver. 1.12.1 (DADA2, Callahan et al., 2016) of $R$ by each MiSeq run. The denoising was performed by a method used in a previous studies (Tsuji, Maruyama, et al., 2020; Tsuji, Miya, et al., 2020; Tsuji, Shibata, et al., 2020). Briefly, first, the "dada2::filterAndTrim" function was used for the removal of excess sequences (primers and random hexamers), and the trimming and filtering of the forward and reverse reads. The error model was trained using the "dada2::learnErrors" function and used to identify and correct indel-mutations and substitutions of dereplicated passed sequences using the "dada2::dada" function. The paired reads were merged, and then the sequences were assigned to each standard DNA sequence or each unknown haplotype using the function "dada2::assignSpecies." The parameters that control the stringency of the denoising step in DADA2, OMEGA_A, and OMEGA_C, were set at the default values. To convert the number of sequence reads to the number of DNA copies, a linear regression line between sequence reads of standard DNAs and their known DNA copies was obtained as a standard line for each sample using the "Im" function in $R$ (the intercept was set to 0 ). For each sample, the number of DNA copies of detected unknown haplotypes was calculated as $\mathrm{Nc}=\mathrm{MRs} / \mathrm{S}$ using the obtained standard line (MRs $=$ the number of MiSeq sequence reads, $S=$ the regression slope of the samplespecific standard line). The Nc values are hereafter referred to as "estimated copy numbers." The detailed workflow of qMiSeq is described in the Ushio et al. (2018) and Tsuji, Shibata, et al. (2020).

## 2.5 | Data screening

To remove chimeric sequences derived from the standard DNAs, the detected haplotypes containing specific sequences to the standard DNAs (std1, 5'-ACA TAT AGT AGT GAG AAC CCA CCA ACT GAT GAT-3'; std2, 5'-GGG ATA ACC ATC CCC TAT ATG GTT TAG TAC ATA- $3^{\prime}$; std3, $5^{\prime}$-TGA GAG ACC ACC AAC TGA TTT GTA TAA AGG TAC-3'; Tsuji, Shibata, et al., 2020) were manually searched and removed. Additionally, the accuracy of eDNA-based detection of haplotypes would be further increased by selecting only haplotypes detected from all filtration replicates because the error sequences are generated randomly and accidentally in each sample (Tsuji, Miya, et al., 2020). Thus, only the haplotypes detected in all three filtration replicates and the average number of sequence reads of them were used for subsequent analysis. To further improve the detection accuracy, the haplotypes detected with an average of less than one copy per liter were removed. The core algorithm of DADA2 merges the sequence reads inferred as errors with the most similar sequence to each (see Callahan et al., 2016 for the detailed algorithm of DADA2). Thus, the sequences with many similar sequences that were recognized as errors may be overestimated in the number of reads. To reduce the observation error, if the coefficient of variation
of the observed number of sequence reads among three filtration replicates exceeded one, an outlier was removed and the average value in the other two replicates was used for subsequent analysis. Previous studies have employed this data screening method and obtained results comparable to large-scale capture surveys (96 individuals/site) (Tsuji, Miya, et al., 2020).

In addition, if the number of detected haplotypes was below the minimum number of that in the previous study (five haplotypes; Iguchi et al., 2002), the sample was excluded from subsequent analyses because they were suspected of having false negative haplotypes due to low eDNA concentrations related to low fish abundances. This is based on the fact that the landlocked Ayu population in Lake Biwa generally migrates to rivers in swarms and is known to have high genetic diversity (Iguchi et al., 2002; Tsuji, Miya, et al., 2020; Tsuji, Shibata, et al., 2020). Hence a large number of haplotypes are expected to be detected in each sample. Furthermore, in the autumn survey in 2018, the water level in all rivers was higher than the normal water levels due to a typhoon that passed through a few days before the survey. The river levels were elevated relative to before the typhoon passed by 1-5 cm for 2-US, 5-AM, and 7-CN; 6-10 cm $6-\mathrm{SO} ; 11-15 \mathrm{~cm}, 3-\mathrm{IN}$ and $8-\mathrm{IS} ; 16-20 \mathrm{~cm}, 4-\mathrm{SR}$ and $11-\mathrm{WN} ;>20 \mathrm{~cm}$, $1-\mathrm{EC}$ and $9-\mathrm{AD}$; and not available for $10-\mathrm{WD}$ (data provided from Flood Management Office of Shiga Prefecture). As autumn migrants have significantly larger biomass than that of spring migrants, it is assumed that a greater variety of haplotypes will be detected from autumn migrants (cf. Azuma, 1970, 1973a, 1973b, 1973c) Nevertheless, in 1-EC and 9-AD, where the water level was particularly high, the number of detected haplotypes were obviously lower than in 2018 spring (see Section 3; Table 1). Thus, the decrease in the
number of detected haplotypes was considered to be more strongly influenced by dilution due to rising water levels than other rivers (Jane et al., 2015). For that reason, 1-EC 2018-autumn and 9-AD 2018-autumn were excluded from subsequent analyses.

## 2.6 | Population genetic analysis

As an indicator to represent intraspecific diversity, haplotype diversity ( $h$; Nei \& Tajima, 1981) was estimated according to the following formula: $h=n\left(1-\sum x_{i}^{2}\right) /(n-1)$, where $n$ is the estimated copy number of each detected haplotype, $x_{i}$ is the frequency of the estimated copy number of haplotype $i$ relative to the total copy number. Furthermore, the pairwise fixation index values ( $F_{\mathrm{ST}}$ ) were estimated using Arlequin ver. 3.5.2.2 (Excoffier \& Lischer, 2010) based on the estimated copy number of each detected haplotype to compare the genetic structure among different sample groups (i.e., between migrant groups for each river and between geographic samples of each migrant group). For the estimation of $F_{S T}$ using the Arlequin, genetic distances were calculated based on uncorrected $p$-distance, and the number of permutations for significance testing was set at 100 (all other settings were default; Slatkin, 1991, 1995; Arlequin user manual, http://cmpg.unibe.ch/software/arlequin35/man/Arlequin35.pdf).

## 2.7 | Statistical analysis

R version 4.1.1 software (R Core Team, 2019) was used for all statistical analyses. The significance level was set at $\alpha=0.05$. To

TABLE 1 Number of detected haplotypes and estimated haplotype diversity ( $h$ ) for each sample

| River | Basin area ( $\mathrm{km}^{2}$ ) | Number of detected haplotypes |  |  |  | Haplotype diversity (h) |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 2018 |  | 2019 |  | 2018 |  | 2019 |  |
|  |  | Spring | Autumn | Spring | Autumn | Spring | Autumn | Spring | Autumn |
| 1-EC | 204.0 | 22 | 13* | 26 | 47 | 0.83 | NA | 0.87 | 0.88 |
| 2-US | 69.7 | $1^{*}$ | 40 | 7 | 62 | NA | 0.87 | 0.78 | 0.89 |
| $3-\mathrm{IN}$ | 101.0 | 43 | 131 | 16 | 81 | 0.89 | 0.90 | 0.84 | 0.90 |
| 4-SR | 65.0 | 46 | 141 | 14 | 95 | 0.89 | 0.91 | 0.75 | 0.92 |
| 5-AM | 109.0 | 9 | 89 | $1 *$ | 57 | 0.82 | 0.89 | NA | 0.90 |
| 6-SO | 19.6 | 39 | 135 | 38 | 56 | 0.88 | 0.91 | 0.91 | 0.88 |
| 7-CN | 51.0 | 84 | 136 | 71 | 85 | 0.89 | 0.91 | 0.92 | 0.90 |
| 8-IS | 59.0 | 53 | 156 | NA | NA | 0.89 | 0.91 | NA | NA |
| 9-AD | 306.0 | 91 | 8* | 53 | 68 | 0.89 | NA | 0.77 | 0.89 |
| 10-WD | 10.0 | 8 | 31 | 9 | 50 | 0.69 | 0.86 | 0.77 | 0.89 |
| 11-WN | 16.0 | 9 | 80 | $1^{*}$ | NA | 0.72 | 0.89 | NA | NA |
|  |  |  |  |  | Average | 0.84 | 0.89 | 0.83 | 0.89 |
|  |  |  |  |  | SD | 0.07 | 0.02 | 0.07 | 0.01 |

Note: The data with an asterisk were excluded from analysis because of the presence of suspected false negative haplotypes due to the low eDNA concentration. NA indicates cases where the water sampling could not be performed due to river construction. The data with asterisks were omitted in the analysis because of the little number of detected haplotypes.
Abbreviation: eDNA, environmental DNA.
confirm the validity of the estimated haplotype diversity using eDNA analysis, the mean values of the estimated $h$ of autumn migrants was compared with that of a previous study based on a capture survey (average $h 0.78$, standard deviation 0.19 ; Iguchi et al., 2002) using the Kruskal-Wallis test (kruskal.test function). The difference in the number of detected haplotypes and the estimated haplotype diversity was tested between and within the migrant groups using the Wilcoxon signed-rank test (exactRankTests package ver. 0.8-32). To account for multiplicity of tests (i.e., each consists of two tests), the Bonferroni correction was applied and the minimum level of significance was set at $p<0.025$. In addition, since the basin area of each river may affect the carrying capacity of the Ayu population, the influence of the basin area on the number of detected haplotypes was tested using a generalized linear mixed-effects model (GLMM; glmmML package ver. 1.1.1) with a Poisson error distribution. In the GLMM model, the study rivers were treated as random effects. Also, a beta regression model (betareg package ver. 3.1-4) was used to investigate the effects of latitude, longitude and presence/absence of fish stocking on $F_{S T}$ between the two migrant groups in each river.

To investigate the spatial genetic structure within each migrant group, the relationship between $F_{S T}$ values and the geographical distance for all sample pairs in each migrant group was assessed using the Mantel correlation test (ecodist package ver. 2.0.7). As a geographical distance among river mouths, three measurement methods were used (Figure 2; Figure S1): (Figure S1a) the shortest straight line distance through the lake, (Figure S1b) right-handed distance along the lakeshore, and (Figure S1c) the shortest distance along the lakeshore. Additionally, Iguchi et al. (2002) reported a significant association between $F_{S T}$ and the geographical distance (the shortest straight line distance through the lake) based on samples from five rivers located in the west side of the lake (U River [ $135.99^{\circ} \mathrm{E}, 35.26^{\circ} \mathrm{N}$ ], 9-AD, 7-CN, Oura River [136.12 ${ }^{\circ} \mathrm{E}, 35.48^{\circ} \mathrm{N}$ ], and $6-\mathrm{SO}$ ). So, as the fourth method, we examined the relationship between $F_{\text {ST }}$ values and the shortest straight line distance through the lake between rivers with (D) dividing Lake Biwa into east and west. Furthermore, principal coordinate analysis (PCoA) based on the $F_{\text {ST }}$ matrix was performed to investigate the genetic structure between two migrant groups and within each migrant group (cmdscale function). To maximize the detection of patterns, adjustment of the significance level for multiple tests was not applied.

To examine the relationship between the geographical location of each river and the spawning timing of Ayu, the probability of spawning at each river was calculated every 2 weeks from the fourth week of August to the second week of November. For the calculations of the spawning probability, the results of 19year surveys of Ayu spawning in Lake Biwa reported by the Shiga Prefectural Fisheries Experiment Station (see Appendix S1 for details, Table S3).

## 3 | RESULTS

## 3.1 | MiSeq sequencing and data screening

The qMiSeq analysis with paired-end sequencing of 153 total libraries (132 real samples, 12 filter negative controls, and 9 PCR negative controls) was successfully performed (total five MiSeq runs, Q30 $>96 \%$ ). $R^{2}$ values of the sample-specific standard line ranged from 0.88 to 1.00 , and the median was 0.94 (Table S4). A total of 695 haplotypes were detected throughout the 2-year survey at 11 rivers. Nucleotide substitution occurred at 63 sites out of 166 bp sequences. The number of sequence reads of detected haplotypes was converted to copy number using the coefficient value of the sample-specific regression line (Table S4). After screening data, 265 haplotypes which were detected from all filtration replicates in each river in each season with one or more copies per liter were finally selected (Table S5).

As a general trend, the number of detected haplotypes for each river ranged from one to 156 (Table 1), with the majority of haplotypes shared among samples from different rivers and/or migrant groups. The data of 2-US (2018-spring), 5-AM (2019-spring) and 11WN (2019-spring) in which only one haplotype was detected were excluded from subsequent analyses because they were suspected of having many false negative haplotypes (Table 1). Additionally, 1-EC (2018-autumn) and 9-AD (2018-autumn) were also excluded from subsequent analyses because they were suspected to have many false negative haplotypes due to high flow (see Section 2; Table 1).

## 3.2 | The number of detected haplotypes and haplotype diversity

The number of detected haplotypes significantly differed between spring and autumn migrants (Wilcoxon signed-rank test, $p<0.01$ in both years), and more haplotypes were detected from the autumn migrants than the spring migrants in both years. On the other hand, there was no significant difference in the number of detected haplotypes between the different year samples of the same migrant group (Wilcoxon signed-rank test, $p=0.125$ for spring, $p=0.078$ for autumn). The GLMM results showed that the number of detected haplotypes was significantly increased with basin area ( $p<0.001$ ).

Haplotype diversity ( $h$ ) of each sample ranged from 0.69 in 10-WD (2018-spring) to 0.92 in $7-\mathrm{CN}$ (2019-spring) for the spring samples and 0.86 in 10-WD (2018-autumn) to 0.92 in 4-SR (2019-autumn) for the autumn samples (Table 1). There was no significant difference between the mean estimated haplotype diversity of autumn migrants based on eDNA analysis and that in previous studies based on capture survey by Iguchi et al. (2002) (KruskalWallis test, $p=0.75)$. The $h$ values significantly differed between the spring and autumn migrants only in 2018 (Wilcoxon signed-rank test, $p<0.01$ for 2018, $p=0.078$ for 2019), but the autumn migrants
showed higher $h$ than that of spring migrants in almost all rivers in both years ( $8 / 8$ rivers for 2018, $6 / 8$ rivers for 2019, except for NA). On the other hand, there was no significant difference in $h$ between the different year samples of the same migrant group (Wilcoxon signed-rank test, $p=0.66$ for spring, $p=0.66$ for autumn).

## 3.3 | Spatio-temporal patterns of genetic differences between the two migrant groups

In each river, the $F_{S T}$ values between two migrant groups varied from 0.0014 (7-CN 2018) to 0.030 (10-WD 2019) (Figure 3; Table S6). All of the calculated $F_{\text {ST }}$ values were low, yet they differed significantly from zero ( $p<0.001$ ). The beta regression model showed a significant negative effect of latitude and the presence of fish stocking on the $F_{S T}$ values between migrant groups in each river (coefficient estimates, -3.16 and $-1.04 ; p<0.05$ and $p<0.01$, respectively).

In both years, there was no significant association between $F_{\text {ST }}$ within each migrant group and three measurements of geographical distance between the rivers when all rivers are taken into account (Mantel test, $p>0.05$; Table 2; Figure 4). On the other hand, when Lake Biwa was divided into east and west sides and analyzed separately, a significant association between $F_{\text {ST }}$ value and the shortest straight line distance was found in the rivers on the west side for spring migrants in both years ( $p<0.05$; Table 2; Figure 4). Although not statistically significant, the east side on both groups in 2018 and the west side on autumn groups in 2019 also showed high Mantel statistic $r$ (0.543-0.943).


FIGURE 3 The relationship between the geographical coordinates of the sampling sites and the genetic distance ( $F_{\mathrm{ST}}$ ) between two migrant groups in each river. The size and color of the circles indicate the magnitude of $F_{S T}$ and the presence ( + ) or absence (-) of fish stoking, respectively. $F_{S T}$ was estimated by comparing different migrant groups in the same year for each river. The data for 2 years (if available) were shown as overlapping circles.

When the $F_{S T}$ values of all sample pairs for each year were ordinated using a PCoA, spring migrants samples were spread widely, whereas autumn migrants samples were clustered in almost the same position, except for 10-WD (2018) (Figure 5a,b). On the other hand, when the $F_{S T}$ values among 2-year samples for each migrant group were ordinated using a PCoA, plots of each river did not tend to cluster closely together (Figure 5c,d).

The probability of Ayu spawning at each river every 2 weeks inferred based on 19-year egg monitoring data (Table S3) showed that Ayu starts to spawn earlier and continues until later in the spawning season in rivers located at higher latitudes (Figure S2).

## 4 | DISCUSSION

Through the two sampling years, we successfully detected and quantified a large number of Ayu haplotypes (a total of 265) from 11 rivers by using eDNA analysis. Iguchi et al. (2002), the only previous study that investigated genetic variation between and within the spring and autumn migrant groups, analyzed 15 specimens from each of five rivers, two of which were used for comparisons between the migrant groups. Iguchi et al. (2002) detected 38 haplotypes in a total of 120 specimens, which is significantly less than our findings. The application of eDNA analysis in our study allowed us to conduct a larger-scale and more extensive inquiry, which Iguchi et al. (2002) had identified as a potential need. Furthermore, using the accurate estimated standard line for each sample, we quantitatively examined haplotype frequency and genetic diversity and successfully obtained reasonable $h$ values that were consistent with the previous study based on the capture survey. As also suggested by Tsuji, Shibata, et al. (2020), the quantitative assessment of genetic diversity based on eDNA analysis enables the application of traditional population genetics tools/indicators. Because the statistical significance of those indicators is dependent on the sample size (in this case, the number of copies per liter), this point should be noticed and examined in the next section.

## 4.1 | Genetic differences between migrant groups

Our $F_{\text {ST }}$ assessment using eDNA revealed genetic differences between spring and autumn migrant groups in respective rivers, although the estimated $F_{S T}$ values were generally small (0.001-0.019 for 2018, 0.003-0.030 for 2019). Additionally, autumn migrants possessed a larger number of haplotypes and showed higher haplotype diversity than spring migrants. This seems reasonable because the former have a much greater abundance than the latter (cf. Azuma, 1970, 1973a, 1973b, 1973c) and supports that spring and autumn migrants are derived from groups with some degree of genetic differences. The small but significant genetic differences between the two migrant groups recognized in this study are consistent with the findings by the capture-based method of Iguchi et al. (2002), which reported $0.063-0.342$ of $F_{S T}$ for two river samples. The

TABLE 2 Results of Mantel tests for associations between the genetic distance ( $F_{\text {ST }}$ ) and the four measurements of geographical distance between study rivers (Figure S1)

| Migrant group | Geographical distance | Year | River | Mantel statistic r | $p$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Spring | The shortest straight line distance through the lake | 2018 | All | 0.263 | 0.15 |
|  |  |  | West side | 0.650 | 0.02 |
|  |  |  | East side | 0.543 | 0.29 |
|  |  | 2019 | All | 0.002 | 0.48 |
|  |  |  | West side | 0.943 | 0.04 |
|  |  |  | East side | -0.143 | 0.52 |
|  | Right-handed distance along the lakeshore from$11-\mathrm{WN}$ | 2018 | All | 0.074 | 0.27 |
|  |  | 2019 | All | 0.176 | 0.13 |
|  | The shortest distance along the lakeshore | 2018 | All | 0.017 | 0.36 |
|  |  | 2019 | All | -0.077 | 0.64 |
| Autumn | The shortest straight line distance through the lake | 2018 | All | -0.071 | 0.54 |
|  |  |  | West side | 0.281 | 0.32 |
|  |  |  | East side | 0.943 | 0.08 |
|  |  | 2019 | All | 0.113 | 0.30 |
|  |  |  | West side | 0.618 | 0.17 |
|  |  |  | East side | -0.073 | 0.50 |
|  | Right-handed distance along the lakeshore from11-WN | 2018 | All | 0.117 | 0.23 |
|  |  | 2019 | All | -0.206 | 0.88 |
|  | The shortest distance along the lakeshore | 2018 | All | 0.003 | 0.38 |
|  |  | 2019 | All | -0.197 | 0.89 |

Note: Bold letters indicate the statistical significances $(p<0.05)$ and the higher Mantel statistic $r(r>0.5)$.
degree of freedom was not determinable for the population statistics calculated from eDNA data. However, at least in the present case, the statistics can be regarded as those based on a large sample size because the population size in each river was sufficiently large (at least hundreds or thousands individuals; cf. Sakai, 2010, 2011) and eDNA could have captured haplotypes from the whole population. Also, a simulation analysis by scaling down the sample size of each to 2000, 1000, and 500 confirmed that the $F_{S T}$ values were still significant in most cases even for the smallest $F_{\text {ST }}$ between spring and autumn migrants (7-CN; Table S7).

## 4.2 | Spatial genetic structure between migrant groups

$F_{S T}$ values between spring and autumn migrants in each river were negatively related to latitude and the presence of fish stocking, suggesting that spatial and anthropogenic factors influence the magnitude of genetic differences between the migrant groups. Of these, the latitude of rivers may influence the degree of overlap in the spawning timing between the two migrant groups because Ayu starts to spawn earlier and continues until later in the spawning season in rivers located at higher latitudes. It has been also known that autumn migrants begin spawning earlier than spring migrants and that some of them can spawn again roughly 20 days following the first spawning (Shimadu, 1950). As a result, autumn migrants in
high-latitude rivers with longer spawning periods are more likely to spawn multiple times. Such multiple spawning by the autumn migrants in higher latitude rivers will increase the chances of simultaneous spawning with spring migrants later in the spawning season, possibly reflecting the decreased $F_{\text {ST }}$ value between migrant groups when assuming some homing migration (but see the next section).

The lower $F_{\text {ST }}$ value between the migrant groups in the rivers where fish stocking has been conducted, on the other hand, could be explained by the admixture of autumn migrant fish into the spring migrant group by the stocking of the latter in the spring season. In five rivers (9-AD, 8-IS, 4-SR, 3-IN, 1-EC), the juveniles of Ayu caught in Lake Biwa (i.e., not spring migrants) have been stocked into the middle and/ or upper reaches every year (lguchi et al., 2002), leading to cohabitation of the two groups during the spring and early summer. Although the proportion and survival until the spawning season of the stocked fish are unclear, the stocked fish may simultaneously spawn with the spring migrants. The overlap in spawning timing and location in this situation could cause the lower $F_{\text {ST }}$ between the migrant groups when evaluated by eDNA or even by the conventional capture-based method.

## 4.3 | Spatial and temporal genetic structure within migrant group

When all rivers were considered, no isolation-by-distance patterns ( $F_{\text {ST }}$ vs. three types of geographic distance) were observed in any


FIGURE 4 Relationship between the geographical distance (the shortest straight line distance through the lake) and the genetic distance $\left(F_{\mathrm{ST}}\right)$ among river samples for each migrant group. The green circle and purple triangle indicate data for 2018 and 2019 , respectively. All, west side and east side indicate the results for data from all rivers studied, rivers on the western side and rivers on the eastern side of Lake Biwa, respectively. The solid and dashed lines indicate statistically significant relationships and the relationships that were not statistically significant but have a higher Mantel statistic $r$ value ( $r>0.5$ ), respectively.
migrant group. Whereas, when Lake Biwa was divided into east and west sides, significant associations between $F_{S T}$ and geographical distance between rivers was observed in spring migrant samples in both years from the west-side rivers. Additionally, higher Mantel statistic $r$ values ( $r>0.5$ ) were observed in three data sets (including autumn migrant samples), although not statistically significant. Iguchi et al. (2002) also reported a significant association between $F_{S T}$ and the geographical distance based on samples from five rivers located in the west side of the lake. These findings point to the prevalence of weak isolation by distance in both migrant groups. On the other hand, isolation by distance was not observed when using the shortest straight line distance between pairs of all the rivers, suggesting that Ayu tends to migrate along the lakeshore instead of taking the route across the center of the lake when the fish search the rivers to enter.

The degree of genetic differentiation among river samples was clearly greater in spring migrants than in autumn migrants. This difference in the genetic structure between the two migrant groups
may be explained by the timing of spawning and hatching and by seasonal changes in gyres (i.e., circulating flow of lake water) in Lake Biwa. As previously stated, the landlocked Ayu in Lake Biwa spawns from the end of August to the beginning of November in the lower sections of inlet rivers, and autumn migrants tend to spawn earlier than spring migrants (Azuma, 1973b; Tsukamoto et al., 1987). The eggs hatch in about 2-3 weeks, and the larval fish immediately flow down to the lake (Azuma, 1970). In Lake Biwa, gyres occur from early summer to autumn, and the largest gyre, called the first gyre, reaches a maximum flow velocity of around $20 \mathrm{~cm} / \mathrm{s}$ in August and September (Figure 2, Akitomo et al., 2004). Therefore, the larvae produced by autumn migrants which flow down to the lake earlier (early to middle September) will likely spread more widely across Lake Biwa by still strong gyres. On the other hand, the larvae which were spawned later mainly by spring migrants with some autumn migrants may tend to stay at the river mouth or littoral areas near their home rivers because of the weakened influences of the gyre in autumn. We here propose a hypothesis that the difference in genetic

FIGURE 5 Results of the principal coordinates analysis (PCoA) on the genetic distance ( $F_{\text {ST }}$ ) between samples for each year ( $a$ and $b$ panels) and between samples for each migrant group (c and d panels). The shapes of each plot indicate each river and are common to all panels. In the 2018 and 2019 panels, orange and blue indicate the data of spring migrants and autumn migrants, respectively. In spring and autumn panels, green and purple indicate data for 2018 and 2019, respectively. In spring and autumn panels, the plots of the same rivers are connected by lines.

structure between the migrant groups is caused by the differences in the extent of diffusion of the larval fish under seasonal fluctuation of gyres.

An alternative explanation for the difference in genetic structure between the migrant groups would be a homing migration habit presumably enhanced in the spring migrant group. However, this explanation was not supported by the comparisons between 2year samples. The PCoA did not show any tendency for different year samples from the same rivers to be plotted close together. This result indicates that the genetic structure changed between the 2 years and does not support simple homing migration. There are two nonexclusive possibilities to explain this result. One is that the spring migrants are shuffled among rivers every year. The tendency of isolation by distance detected in this study suggests that such shuffling seems to be limited only between neighboring rivers. However, because Ayu has been inferred to migrate into a nearby river with relatively higher water levels at the time of their migration (Takahashi \& Azuma, 2016), fish born in different rivers can be mixed with each other. Another explanation is based on the "switching hypothesis" proposed by Tsukamoto et al. (1987). This hypothesis was built based on a growth analysis using otoliths and the spawning timing of each migrant group, and involves the change of the migration type between generations (Figure S3). That is, early-born individuals (expected as the progeny of autumn migrants) migrate to the river in the next spring, whereas late-born individuals (expected as the progeny of spring migrants) migrate to the river in the next autumn. Since the autumn migrant group is always far more numerous than the spring migrant group, complete switching of the migration pattern between generations is unlikely. However, even if only the earliest
spawners of autumn migrants and the latest spawners of spring migrants switch with each other, the genetic structure would change between 2018 and 2019. To test these hypotheses, examination over several generations is necessary.

## 5 | CONCLUSION

Our results suggested that there were genetic differences between two migrant groups of the landlocked Ayu population inhabiting Lake Biwa, and the $F_{S T}$ values were negatively related to latitude and fish stocking. Additionally, there was spatial and temporal genetic structure between and within migrant groups. Although further explicit tests are necessary, genetic differentiation between migrant groups and geographical population structure were hypothesized to be influenced by differences in spawning timing between populations and lake currents. Further studies on ecological and genetic factors determining migration timing would give better understanding toward the evolution and maintenance mechanisms of the life history divergence in the Ayu population of Lake Biwa. Finally, as this study has demonstrated, the eDNA approach can be effective for conducting large-scale genetic structural investigations and contribute as a springboard for further detailed population genetic studies.

## AUTHOR CONTRIBUTIONS

Satsuki Tsuji conceived and designed research. Satsuki Tsuji, Naoki Shibata, and Hayato Sawada performed field surveys and molecular analysis. Satsuki Tsuji and Katsutoshi Watanabe performed data analysis. Satsuki Tsuji and Katsutoshi Watanabe wrote the
early draft and it was completed with significant input from all authors.

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## CONFLICT OF INTEREST

None declared.

## DATA AVAILABILITY STATEMENT

Full details of the results for each experiment of the present study are available in the Appendix and supporting information (Tables S1-S7). All raw sequences obtained in the qMiFish analysis were deposited in the DDBJ Sequence Read Archive (accession number: DRA013835).

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## SUPPORTING INFORMATION

Additional supporting information can be found online in the Supporting Information section at the end of this article.

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