

New scheme for pricing Bermudan options under stochastic volatility model

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January 30, 2012

Abstract

The author considers stochastic volatility models and introduces a new scheme for pricing Bermudan options under stochastic volatility models. His approach is the asymptotic expansion method which is based on Malliavin calculus.

1 Introduction

The valuation of Bermudan options is very important problem in option pricing theory. The values of Bermudan options in stochastic volatility models are calculated with the regression method developed by Longstaff and Schwartz [3]. This method is not suitable for parallel computing.

In this paper, we introduce a new scheme for pricing Bermudan options. This scheme is very universal and can be applied to problems we can not develop recombining trees. For example, we can apply to evaluations of derivatives under SV models.

Our scheme has two keys. One is to derive an approximate formula of the joint distribution function of stochastic processes using the asymptotic expansion method. The other is to develop recombining tree with the idea of binning [2] using the approximate joint distribution function. Using the recombining tree, we evaluate derivatives like Bermudan options under stochastic volatility models. Our scheme is suitable for parallel computing.

The structure of this paper is as follows. The next section reviews the stochastic volatility models which are widely accepted in financial industry and applies the asymptotic expansion method to the model. The 3rd section describes how to derive our approximate formula of the joint distribution functions with the asymptotic expansion method, while the following section derives the

joint distribution function of SABR model. The 5th section presents numerical results of our new scheme. The final section concludes.

2 Stochastic volatility model

2.1 Definition of stochastic volatility model

Let $(\Omega, \mathcal{F}, \mathbb{P}, \{\mathcal{F}_t\}_{0 \leq t \leq T})$ be a complete probability space satisfying the usual hypotheses and $T \in (0, \infty)$ denotes some fixed horizon of economy. Let $(W_1(t), W_2(t))$, $0 \leq t \leq T$, be a 2-dimensional correlated Brownian motion with correlation given by $\rho : [0, T] \rightarrow [-1, 1]$ such that

$$d\langle W_1, W_2 \rangle_t = \rho(t) dt. \quad (1)$$

We consider the following stochastic differential equation for X and Y :

$$dX(t) = B(t, X(t), Y(t)) dW_1(t), \quad (2)$$

$$dY(t) = M(t, Y(t)) dt + D(t, Y(t)) dW_2(t), \quad (3)$$

$$(X(0), Y(0)) = (x_0, y_0) \in \mathbb{R} \times \mathbb{R}, \quad (4)$$

Suppose B , M and D satisfy some regularity conditions.

2.2 Asymptotic expansion of stochastic volatility model

We consider an perturbed stochastic process defined as the following stochastic differential equation:

$$dX^\epsilon(t) = \epsilon B(t, X^\epsilon(t), Y^\epsilon(t)) dW_1(t), \quad (5)$$

$$dY^\epsilon(t) = M(t, Y^\epsilon(t)) dt + \epsilon D(t, Y^\epsilon(t)) dW_2(t), \quad (6)$$

$$(X^\epsilon(0), Y^\epsilon(0)) = (x_0, y_0) \in \mathbb{R} \times \mathbb{R}. \quad (7)$$

We want to calculate an approximate solution of this model by using the asymptotic expansion approach. By results of [5], we have the following lemma.

Lemma 2.1. *$X^\epsilon(t)$ and $Y^\epsilon(t)$ have following approximate solutions as $\epsilon \rightarrow 0$ respectively.*

$$X^\epsilon(T) = \sum_{i=0}^N \epsilon^i X_i(T) / i! + o(\epsilon^N), \quad (8)$$

$$Y^\epsilon(T) = \sum_{i=0}^N \epsilon^i Y_i(T) / i! + o(\epsilon^N), \quad (9)$$

where

$$X_i(T) = \left. \frac{d^i X^\epsilon(T)}{d\epsilon^i} \right|_{\epsilon=0}, \quad (10)$$

$$Y_i(T) = \left. \frac{d^i Y^\epsilon(T)}{d\epsilon^i} \right|_{\epsilon=0}, \quad (11)$$

for $i = 0, 1, \dots, N$.

Here, we can calculate $X_i(T)$ and $Y_i(T)$ analytically. Examples of $Y_i(T)$ are as follows:

$$Y_1^\epsilon(T) = \tilde{M}(T) \int_0^T \tilde{M}(t_1)^{-1} D(t_1, Y_0(t_1)) dW_2(t_1) \quad (12)$$

$$\begin{aligned} Y_2^\epsilon(T) &= \tilde{M}(T) \int_0^T \tilde{M}(t_1)^{-1} Y_1(t_1)^2 M_{y,y}(t_1, Y_0(t_1)) dt_1, \\ &+ 2 \tilde{M}(T) \int_0^T \tilde{M}(t_1)^{-1} Y_1(t_1) D_y(t_1, Y_0(t_1)) dW_2(t_1) \end{aligned} \quad (13)$$

$$\begin{aligned} Y_3^\epsilon(T) &= \tilde{M}(T) \int_0^T \tilde{M}(t_1)^{-1} Y_1(t_1)^3 M_{y,y,y}(t_1, Y_0(t_1)) dt_1, \\ &+ 3 \tilde{M}(T) \int_0^T \tilde{M}(t_1)^{-1} Y_1(t_1) Y_2(t_1) M_{y,y}(t_1, Y_0(t_1)) dt_1, \\ &+ 3 \tilde{M}(T) \int_0^T \tilde{M}(t_1)^{-1} Y_1(t_1)^2 D_{y,y}(t_1, Y_0(t_1)) dW_2(t_1), \\ &+ 3 \tilde{M}(T) \int_0^T \tilde{M}(t_1)^{-1} Y_2(t_1) D_y(t_1, Y_0(t_1)) dW_2(t_1) \end{aligned} \quad (14)$$

where

$$\tilde{M}(T) = \exp \left(\int_0^T M_y(t_0, Y_0(t_0)) dt_0 \right). \quad (15)$$

$$(16)$$

And examples of $X_i(T)$ are as follows:

$$X_0^\epsilon(T) = x_0 \quad (17)$$

$$X_1^\epsilon(T) = \int_0^T B(t_0, Y_0(t_0), X_0(t_0)) dW_1(t_0) \quad (18)$$

$$X_2^\epsilon(T) = 2 \int_0^T X_1(t_0) B_y(t_0, Y_0(t_0), X_0(t_0)) dW_1(t_0) \quad (19)$$

$$+ 2 \int_0^T Y_1(t_0) B_x(t_0, Y_0(t_0), X_0(t_0)) dW_1(t_0) \quad (20)$$

$$X_3^\epsilon(T) = 3 \int_0^T X_1(t_0)^2 B_{y,y}(t_0, Y_0(t_0), X_0(t_0)) dW_1(t_0) \quad (21)$$

$$+ 3 \int_0^T X_2(t_0) B_y(t_0, Y_0(t_0), X_0(t_0)) dW_1(t_0) \quad (22)$$

$$+ 6 \int_0^T X_1(t_0) Y_1(t_0) B_{x,y}(t_0, Y_0(t_0), X_0(t_0)) dW_1(t_0) \quad (23)$$

$$+ 3 \int_0^T Y_1(t_0)^2 B_{x,x}(t_0, Y_0(t_0), X_0(t_0)) dW_1(t_0) \quad (24)$$

$$+ 3 \int_0^T Y_2(t_0) B_x(t_0, Y_0(t_0), X_0(t_0)) dW_1(t_0) \quad (25)$$

3 Approximation formula of the joint distribution function

We have to calculate conditional expectations to derive an approximate formula of the joint distribution function. The next theorem is very useful to calculate conditional expectations.

Theorem 3.1. *Let $f \in \mathbb{L}^2(\mathbb{T}^n)$ for $n \geq 1$, $q_1^j \in \mathbb{L}(\mathbb{T})$ for $1 \leq j \leq m$. Let $\{W_i\}_{i=1,\dots,n}$ be an n -dimensional correlated Brownian motion and $\{Z_i\}_{i=1,\dots,m}$ be an m -dimensional correlated Brownian motion. We denote $(t_1, t_2, \dots, t_n)$ by $(\mathbf{t})$.*

$$\begin{aligned} & \mathbb{E} \left[\int_0^T \int_0^{t_1} \dots \int_0^{t_{n-1}} f(\mathbf{t}) dW_n(t_n) \dots dW_2(t_2) dW_1(t_1) \mid \right. \\ & \quad \left. \left\{ \int_0^T q_1^1(t) dZ_1(t), \dots, \int_0^T q_1^m(t) dZ_m(t) \right\} = \{c_1, \dots, c_m\} \right] \\ &= \int_0^T \int_0^{t_1} \dots \int_0^{t_{n-1}} f(\mathbf{t}) \hat{H}_n(\mu(\mathbf{t}), \Sigma(\mathbf{t})) dt_n \dots dt_2 dt_1, \end{aligned} \quad (26)$$

where

$$d \langle W_i, Z_j \rangle = \rho_{i,j} dt, \quad (27)$$

$$\Sigma_c = \left\{ \int_0^T q_i(t) q_j(t) dt \right\}_{i,j=1,\dots,m}, \quad (28)$$

$$\tilde{\Sigma}(t) = \{\rho_{i,j} q_j(t_i)\}_{i=1,\dots,n, j=1,\dots,m}, \quad (29)$$

$$\mu(t) = \Sigma_c^{-1} \tilde{\Sigma}(t), \quad (30)$$

$$\Sigma(t) = -\tilde{\Sigma}(t) \Sigma_c^{-1} \tilde{\Sigma}(t), \quad (31)$$

$$m(\xi; \mu(t), \Sigma(t)) = \exp(\mu(t)^T \xi + 1/2 \xi^T \Sigma(t) \xi), \quad (32)$$

$$\tilde{H}_n(\mu(t), \Sigma(t)) = \frac{d^n m(\xi; \mu(t), \Sigma(t))}{d\xi_1 \cdots d\xi_n} \Big|_{\xi=0}. \quad (33)$$

Let $X_G^\epsilon(T) = (X^\epsilon(T) - X_0(T))/\epsilon$ and $Y_G^\epsilon(T) = (\sigma^\epsilon(T) - \sigma_0(T))/\epsilon$. We want to derive the joint distribution function of $X_G^\epsilon(T)$ and $Y_G^\epsilon(T)$. Let $\varphi_{X_G, Y_G} : \mathbb{R} \times \mathbb{R} \rightarrow \mathbb{R}$ be a characteristic function of $X_G^\epsilon(T)$ and $Y_G^\epsilon(T)$.

Proposition 3.1. $\varphi_{X,Y}$ has an approximate expression as follows:

$$\varphi_{X_G, Y_G}(\xi_1, \xi_2) = \sum_{i=0}^N \frac{\epsilon^i}{i!} \frac{d^i \mathbb{E} [\exp(\sqrt{-1} \xi_1 X^\epsilon(T) + \sqrt{-1} \xi_2 Y^\epsilon(T))]}{d\epsilon^i} \Big|_{\epsilon=0} + o(\epsilon^N) \quad (34)$$

for $(\xi_1, \xi_2) \in \mathbb{R} \times \mathbb{R}$.

In case that $N = 2$,

$$\begin{aligned} \varphi_{X_G, Y_G}(\xi_1, \xi_2) &= \mathbb{E}[N(T)] + \frac{\sqrt{-1}}{2} \mathbb{E}[(\xi_1 X_2(T) + \xi_2 Y_2(T)) N(T)] \\ &\quad + \frac{\sqrt{-1}\epsilon}{6} \mathbb{E}[(\xi_1 X_3(T) + \xi_2 Y_3(T)) N(T)] \\ &\quad - \frac{\epsilon^2}{8} \mathbb{E}[(\xi_1 X_2(T) + \xi_2 Y_2(T))^2 N(T)] + o(\epsilon^2) \end{aligned} \quad (35)$$

where

$$N(T) = \exp(\sqrt{-1} \xi_1 X_1(T) + \sqrt{-1} \xi_2 Y_1(T)). \quad (36)$$

By using the inversion formulas of characteristic functions, we get an approximate formula of the joint probability density function of $X_G^\epsilon(T)$ and $Y_G^\epsilon(T)$.

Proposition 3.2. $X_G^\epsilon(T)$ and $Y_G^\epsilon(T)$ have a 3rd order approximate joint prob-

ability density function f_{X_G, Y_G} as follows:

$$\begin{aligned}
 f_{X_G, Y_G}(x, y) &= n(x, y; \Sigma) - \frac{1}{2} \frac{d}{dx} \{ \mathbb{E}^c [X_2(T)] n(x, y; \Sigma) \} - \frac{1}{2} \frac{d}{dy} \{ \mathbb{E}^c [Y_2(T)] n(x, y; \Sigma) \} \\
 &\quad - \frac{1}{6} \frac{d}{dx} \{ \mathbb{E}^c [X_3(T)] n(x, y; \Sigma) \} - \frac{1}{6} \frac{d}{dy} \{ \mathbb{E}^c [Y_3(T)] n(x, y; \Sigma) \} \\
 &\quad + \frac{1}{8} \frac{d^2}{dx^2} \{ \mathbb{E}^c [X_2(T)^2] n(x, y; \Sigma) \} + \frac{1}{8} \frac{d^2}{dy^2} \{ \mathbb{E}^c [Y_2(T)^2] n(x, y; \Sigma) \} \\
 &\quad + \frac{1}{4} \frac{d^2}{dx dy} \{ \mathbb{E}^c [X_2(T) Y_2(T)] n(x, y; \Sigma) \}, \tag{37}
 \end{aligned}$$

for $(x, y) \in \mathbb{R} \times \mathbb{R}$, where

$$\mathbb{E}^c[\cdot] = \mathbb{E}[\cdot | (X_1(T), Y_1(T)) = (x, y)], \tag{38}$$

$$n(x, y; \Sigma) = \frac{1}{2\pi\sqrt{|\Sigma|}} \exp(-[x, y] \Sigma^{-1} [x, y]), \tag{39}$$

$$\Sigma = \begin{bmatrix} \mathbb{E}[X_1(T)^2] & \mathbb{E}[X_1(T) Y_1(T)] \\ \mathbb{E}[X_1(T) Y_1(T)] & \mathbb{E}[Y_1(T)^2] \end{bmatrix}. \tag{40}$$

Then, $X^\epsilon(T)$ and $Y^\epsilon(T)$ have a 3rd order approximate joint distribution function $F_{X, Y}$ as follows:

$$F_{X, Y}(x, y) = \int_0^{x-X_0(T)} \int_0^{y-Y_0(T)} f_{X_G, Y_G}(v, w) dw dv \tag{41}$$

We can calculate conditional expectations in the above lemma by using Theorem 3.1.

4 Pricing Bermudan options

We introduce a new scheme for pricing Bermudan options under stochastic volatility models in this section. In order to clarify the dependency of the variables, we use notations as follows:

$$\begin{aligned}
 F_{X, Y}(x_0, y_0, T, x, y) &= \mathbb{P}(X(T) \leq x, Y(T) \leq y | X(0) = x_0, Y(0) = y_0). \tag{42}
 \end{aligned}$$

$$\begin{aligned}
 \mathbb{P}_{X, Y}(x_0, y_0, T, l_x, u_x, l_y, u_y) &= \mathbb{P}(l_x \leq X(T) \leq u_x, l_y \leq Y(T) \leq u_y | X(0) = x_0, Y(0) = y_0) \tag{43}
 \end{aligned}$$

We approximate $\mathbb{P}_{X, Y}(x_0, y_0, T, l_x, u_x, l_y, u_y)$ using results of Section 3. First, we have an approximate joint distribution function of X and Y by Proposition 3.2. Second, we calculate conditional expectations in the approximate joint distribution function using Theorem 3.1. Then we have an approximate formula of $\mathbb{P}_{X, Y}(x_0, y_0, T, l_x, u_x, l_y, u_y)$.

4.1 Bermudan options

Let $\mathbb{T}$ be $[T_0 = 0, T_1, T_2, \dots, T_n, \infty]$ for $n \geq 1$ and $\mathcal{T}$ be a set of stopping time $\tau : \Omega \rightarrow \mathbb{T}$. We want to calculate a value $V(t)$ that is defined as follows:

$$V(t) = \sup_{\tau \in \mathcal{T}} \mathbb{E} \left[C(\tau, X(\tau), Y(\tau)) \middle| \mathcal{F}_t \right]. \quad (44)$$

We consider this option in this section.

4.2 New scheme

Let $\mathbb{X}$ be $[x_1, x_2, \dots, x_N]$ and $\mathbb{Y}$ be $[y_1, y_2, \dots, y_M]$ for $N \geq 1$ and $M \geq 1$ respectively. We define a_i for $0 \leq i \leq N$ and b_j for $0 \leq j \leq M$ as follows:

$$a_i = \begin{cases} -\infty & i = 0 \\ (x_i + x_{i+1})/2 & i = 1, 2, \dots, N-1, \\ \infty & i = N \end{cases} \quad (45)$$

$$b_i = \begin{cases} -\infty & i = 0 \\ (x_i + x_{i+1})/2 & i = 1, 2, \dots, M-1. \\ \infty & i = M \end{cases} \quad (46)$$

We calculate the value $V(k, i, j)$ of the option at time T_k and $(X(T_k), Y(T_k)) = (x_i, y_j)$ as follows:
when $k = n$,

$$V(k, i, j) = C(T_k, x_i, y_j), \quad (47)$$

otherwise,

$$V(k, i, j) = \max \left(C(T_k, x_i, y_j), \sum_{\tilde{i}=1, \tilde{j}=1}^{N, M} V(k+1, \tilde{i}, \tilde{j}) \mathbb{P}(i, j, k+1, \tilde{i}, \tilde{j}) \right), \quad (48)$$

where

$$\mathbb{P}(i, j, k+1, \tilde{i}, \tilde{j}) = \mathbb{P}(x_i, y_i, T_{k+1} - T_k, a_{\tilde{i}-1}, a_{\tilde{i}}, b_{\tilde{j}-1}, b_{\tilde{j}}). \quad (49)$$

Derivatives are valued in this scheme by the usual backward induction method. Since a direct construction of a multidimensional tree would not lead to recombining nodes, the computational effort would grown exponentially in the number of time steps. However, the computational effort is $n \times N \times M$ in our scheme.

Table 1: Parameter

	x_0	y_0	α	β	ρ	ϵ	r
(i)	100	0.3	0.3	1.0	0.2	1.0	0.01
(ii)	100	0.3	0.3	0.5	0.2	1.0	0.01

5 Numerical result

To test the validity of the new scheme, we consider Bermudan and European put option under the SABR model as follows:

$$dX^\epsilon(t) = \epsilon Y^\epsilon(t) X^\epsilon(t)^\beta dW_1(t), \quad (50)$$

$$dY^\epsilon(t) = \epsilon \alpha Y^\epsilon(t) dW_2(t), \quad (51)$$

$$d\langle X^\epsilon, Y^\epsilon \rangle_t = \rho dt, \quad (52)$$

$$(X^\epsilon(0), Y^\epsilon(0)) = (x_0, y_0) \in \mathbb{R}^+ \times \mathbb{R}^+, \quad (53)$$

$$S(T) = \exp(rT) X^\epsilon(T). \quad (54)$$

Let execution times of Bermudan options be $\mathbb{T} = \{1.0, 2.0, 3.0, 4.0\}$ and the maturity of European option be $T = 4.0$. We calculate following values.

$$\text{Put}_{\text{Eur}} = \mathbb{E} \left[\exp(-rT) (K - S(T))^+ \right], \quad (55)$$

$$\text{Put}_{\text{Ber}} = \sup_{\tau \in \mathcal{T}} \mathbb{E} \left[\exp(-r\tau) (K - S(\tau))^+ \right], \quad (56)$$

where $\mathcal{T}$ is a set of stopping time $\tau : \Omega \rightarrow \mathbb{T}$ and K is strike.

In the test of the new scheme, we set $N = 100$ and $M = 50$, and define x_1 , x_N , y_1 and y_M as follows:

$$x_1 = \mathbb{E}[X^\epsilon(T)] + 5\mathbb{E} \left[(X^\epsilon(T) - X_0(T))^2 \right]^{1/2} \quad (57)$$

$$x_N = \mathbb{E}[X^\epsilon(T)] - 5\mathbb{E} \left[(X^\epsilon(T) - X_0(T))^2 \right]^{1/2} \quad (58)$$

$$y_1 = \mathbb{E}[X^\epsilon(T)] + 5\mathbb{E} \left[(X^\epsilon(T) - X_0(T))^2 \right]^{1/2} \quad (59)$$

$$y_M = \mathbb{E}[Y^\epsilon(T)] - 5\mathbb{E} \left[(X^\epsilon(T) - X_0(T))^2 \right]^{1/2} \quad (60)$$

$$(61)$$

The model parameters used in the test are given in Table 1. We use a 4th order asymptotic expansion for the joint distribution function and an approximate cumulative bivariate normal probabilities[1].

We use values which are calculated in Monte Carlo simulations as benchmarks. In the simulations, we use Ninomiya-Victoir scheme[4] as a discretization scheme with 8 time steps per a year and generate 10^7 paths in each simulation.

Results are in Table 2. We compare our estimations of values by an asymptotic expansion with forth order to the bechmarks.

Table 2: Numerical results

Case	Strike	Value Put _{Ber} (A.E.)	Value Put _{Eur} (A.E.)	Value Put _{Eur} (M.C.)	Imp.Vol. Put _{Eur} (A.E.)	Imp.Vol. Put _{Eur} (M.C.)	Error (bpt)	Prob(ITM)
(i)	50	2.9812	2.7406	2.6002	0.3220	0.3168	51.620	0.17
(i)	55	3.9393	3.6815	3.5629	0.3171	0.3136	35.840	0.22
(i)	60	5.1114	4.8334	4.7367	0.3137	0.3112	24.720	0.27
(i)	65	6.5411	6.2400	6.1324	0.3120	0.3096	23.860	0.31
(i)	70	8.1883	7.8583	7.7558	0.3106	0.3086	20.170	0.36
(i)	75	10.0547	9.6885	9.6075	0.3096	0.3082	14.460	0.41
(i)	80	12.1385	11.7289	11.6842	0.3089	0.3082	7.340	0.46
(i)	85	14.4529	13.9930	13.9778	0.3088	0.3085	2.330	0.50
(i)	90	17.0126	16.4966	16.4789	0.3095	0.3092	2.570	0.54
(i)	95	19.7618	19.1804	19.1749	0.3102	0.3102	0.760	0.58
(i)	100	22.6899	22.0339	22.0528	0.3111	0.3113	-2.530	0.62
(i)	105	25.7867	25.0463	25.0984	0.3120	0.3127	-6.820	0.65
(i)	110	29.0496	28.2181	28.2975	0.3131	0.3141	-10.210	0.68
(i)	115	32.4811	31.5532	31.6370	0.3147	0.3157	-10.640	0.71
(i)	120	36.0408	35.0061	35.1040	0.3162	0.3174	-12.310	0.74
(i)	125	39.7158	38.5677	38.6868	0.3176	0.3191	-14.930	0.76
(i)	130	43.4977	42.2293	42.3748	0.3191	0.3209	-18.250	0.78
(i)	135	47.3857	45.9889	46.1573	0.3206	0.3227	-21.180	0.80
(i)	140	51.3770	49.8482	50.0250	0.3224	0.3246	-22.360	0.81
(i)	145	55.4463	53.7786	53.9694	0.3240	0.3265	-24.320	0.83
(i)	150	59.5876	57.7742	57.9833	0.3257	0.3283	-26.930	0.84
(ii)	90	0.0544	0.0526	0.0548	0.0359	0.0361	-2.310	0.02
(ii)	92	0.0996	0.0944	0.0968	0.0344	0.0346	-1.603	0.04
(ii)	94	0.1857	0.1697	0.1724	0.0331	0.0333	-1.148	0.06
(ii)	96	0.3515	0.3050	0.3077	0.0321	0.0321	-0.773	0.09
(ii)	98	0.6701	0.5431	0.5442	0.0313	0.0313	-0.226	0.16
(ii)	100	1.2559	0.9389	0.9379	0.0308	0.0308	0.164	0.26
(ii)	102	2.2185	1.5514	1.5494	0.0307	0.0306	0.264	0.38
(ii)	104	3.5703	2.4210	2.4205	0.0308	0.0308	0.064	0.52
(ii)	106	5.2187	3.5534	3.5547	0.0314	0.0314	-0.171	0.65
(ii)	108	7.0428	4.9166	4.9182	0.0322	0.0322	-0.229	0.76
(ii)	110	8.9550	6.4566	6.4591	0.0331	0.0332	-0.435	0.84
(ii)	112	10.9065	8.1231	8.1269	0.0342	0.0343	-0.821	0.89
(ii)	114	12.8745	9.8757	9.8807	0.0353	0.0355	-1.380	0.93
(ii)	116	14.8495	11.6856	11.6913	0.0365	0.0367	-2.028	0.95
(ii)	118	16.8276	13.5330	13.5391	0.0377	0.0380	-2.860	0.97
(ii)	120	18.8071	15.4045	15.4111	0.0388	0.0392	-4.078	0.97

A Proof of Theorem 3.1

A.1 preliminaries

Lemma A.1. *Fixed $T \in (0, \infty)$. Let $\mathbb{T} = [0, T]$, μ be the Lebesgue measure, $f_n \in \mathbb{L}^2(\mathbb{T}^n, \sigma(\mathbb{T})^n, \mu^n)$ for $n \geq 1$ and $(W_1, W_2, \dots, W_n)$ be a n -dimensional correlated Brownian motion. We denote by $\mathcal{E}_n$ the set of elementary functions of the form*

$$f(\mathbf{t}) = \sum_{i_1, \dots, i_n=1}^k c_{i_1 \dots i_n} \mathbf{1}_{A_{i_1} \times \dots \times A_{i_n}}(\mathbf{t}) \quad (62)$$

where $A_1, \dots, A_k$ are pairwise-disjoint sets belonging to $\sigma(\mathbb{T})$, and the coefficients $c_{i_1 \dots i_n}$ are zero if any two of the indices $i_1, \dots, i_n$ are equal. Then there exists a sequence $\{f_n^{(l)}\}_{l \in \mathbb{N}} \in \mathcal{E}_n$ such that $f_n^{(l)} \nearrow f_n$ and

$$\begin{aligned} & \mathbb{E} \left[\int_0^T \dots \int_0^T f_n^{(l)}(\mathbf{t}) dW_n(t_n) \dots dW_1(t_1) \middle| \mathcal{G} \right] \rightarrow \\ & \mathbb{E} \left[\int_0^T \dots \int_0^T f_n(\mathbf{t}) dW_n(t_n) \dots dW_1(t_1) \middle| \mathcal{G} \right] (a.s.), \end{aligned} \quad (63)$$

where $\mathcal{G} \subset \sigma(\mathbb{T})$.

A.2 Proof

We use symbols in Lemma A.1. We set $\mathcal{G}$ as follows:

$$\mathcal{G} = \left\{ \left(\int_0^T q_1^1(t) dZ_1(t), \dots, \int_0^T q_1^m(t) dZ_m(t) \right) = (c_1, \dots, c_m) \right\}. \quad (64)$$

Then we have

$$\begin{aligned} & \mathbb{E} \left[\int_0^T \dots \int_0^T f_n^{(l)}(\mathbf{t}) dW_n(t_n) \dots dW_1(t_1) \middle| \mathcal{G} \right] \\ &= \sum_{i_1, \dots, i_n=1}^k c_{i_1 \dots i_n} \mathbb{E} \left[\int_0^T \mathbf{1}_{A_{i_n}}(t) dW_n(t) \dots \int_0^T \mathbf{1}_{A_{i_1}}(t) dW_1(t) \middle| \mathcal{G} \right] \\ &= \int_0^T \dots \int_0^T \sum_{i_1, \dots, i_n=1}^k c_{i_1 \dots i_n} \mathbf{1}_{A_{i_1} \times \dots \times A_{i_n}}(\mathbf{t}) \hat{H}_n(\mu(\mathbf{t}), \Sigma(\mathbf{t})) dt_n \dots dt_1 \\ &= \int_0^T \dots \int_0^T f_n^{(l)}(\mathbf{t}) \hat{H}_n(\mu(\mathbf{t}), \Sigma(\mathbf{t})) dt_n \dots dt_1 \end{aligned} \quad (65)$$

$$\rightarrow \int_0^T \dots \int_0^T f_n(\mathbf{t}) \hat{H}_n(\mu(\mathbf{t}), \Sigma(\mathbf{t})) dt_n \dots dt_1 \quad (66)$$

We define $f_n(t)$ as follows:

$$f_n(t) = 1_{\{t_n \leq \dots \leq t_1\}}(t) f(t), \quad (67)$$

then we have Theorem 3.1.

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